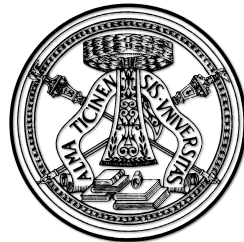


NONPARAMETRIC IDENTIFICATION OF PHARMACOKINETIC MODELS VIA GAUSSIAN PROCESSES

Giuseppe DE NICOLAO

*University of Pavia
Italy*



Outline

- ❖ Bayesian Learning
 - ❖ Population Data Analysis
 - ❖ Signal Estimation
 - ❖ Hyperparameter Estimation: Empirical Bayes and MCMC
 - ❖ Application to PK Data
-

Bayesian Learning

Bayesian Learning

Model: $y_i = f(x_i) + v_i$ $v \sim N(0, \sigma^2 I)$

Problem: estimate $f(\cdot) : \mathbb{R}^d \mapsto \mathbb{R}$ from the
training set $D_n = \{(x_i, y_i), i = 1, \dots, n\}$

Solution: Bayesian approach

Bayes Theorem:

$$p(f | D_n) = \frac{p(D_n | f) p(f) ?}{p(D_n)}$$

Diagram illustrating Bayes Theorem with labels:

- $p(f | D_n)$ is labeled *posterior* (indicated by an arrow).
- $p(D_n | f)$ is labeled *likelihood* (indicated by an arrow).
- $p(f)$ is labeled *prior* (indicated by an arrow).
- The product $p(D_n | f) p(f)$ is circled in red, and a red question mark is placed next to it.

Bayesian Learning

$f(\cdot) \sim$ Gaussian Process (GP)

GP properties:

- easy to compute
- flexibility
- determined by their mean
and autocovariance function $K(x_1, x_2)$

$$K(x_1, x_2) = K(x_1, x_2; \underbrace{\lambda_1^2, \dots, \lambda_r^2}_{\text{hyperparameters}})$$



GCV, OCV, C_p ,
ML, MCMC, ...

Bayesian Learning

$f(\cdot) \sim \text{GP} \longrightarrow p(f \mid D_n)$ is Gaussian

\downarrow if $E[f] = 0$

$$\hat{f}(x) = E[f \mid D_n] = \bar{K}(x)^T \text{Var}[y]^{-1} y$$

where $\bar{K}(x) = [K(x, x_1) \cdots K(x, x_n)]^T$

Remark:

Let $\theta = \text{Var}[y]^{-1} y \longrightarrow$

$$\hat{f}(x) = \sum_{i=1}^n K(x, x_i) \theta_i$$

Radial K



RBF

Bayesian Learning

Tychonov Regularization \longleftrightarrow Bayesian Estimation



$$f_{\gamma}(x) = \arg \min_f \sum_{j=1}^n (y_j - f(x_j))^2 + \underbrace{\gamma \|f\|_K^2}_{\text{"smoothness penalty"}}$$

Solution: $\hat{f}_{\gamma}(x) = \sum_{j=1}^n K(x, x_j) \tilde{\theta}_j$

Remark:

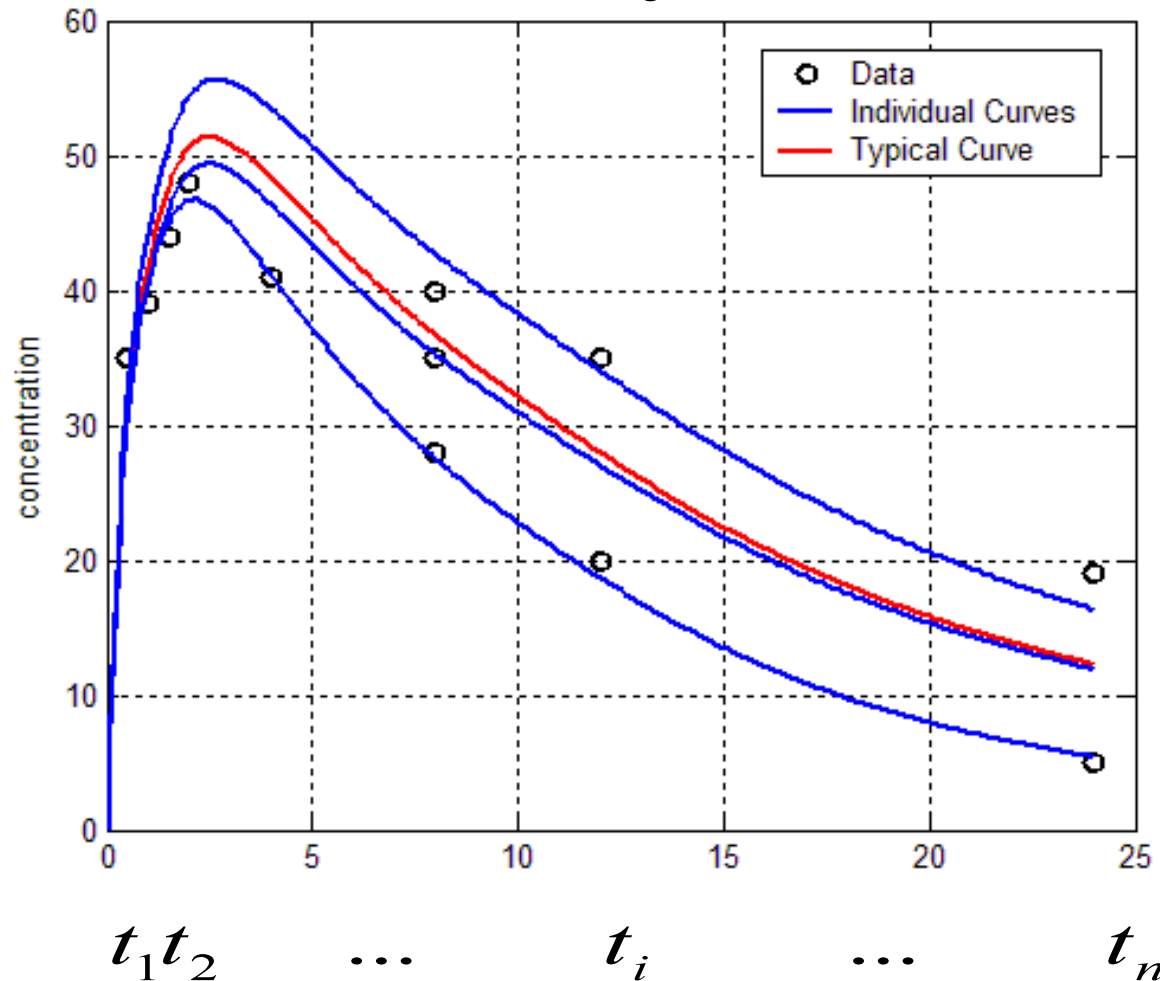
if $\hat{\gamma} = \text{Var}[v] \longrightarrow \tilde{\theta} = \theta \longrightarrow \hat{f}_{\gamma}(x) = \hat{f}(x)$

Bayes estimator

Population Data Analysis

Pharmacokinetic data

N subjects



Noisy
data

$$y_k = x(t_k) + v_k$$

Main goals:

- Typical curve
- Individual curves

Population data analysis: an example

State of the art

- ***Parametric models:***

Example: compartmental models

Drawback: a structural model is needed

- ***Semiparametric models:***

Example: regression splines [*Verotta et al., 1997*]

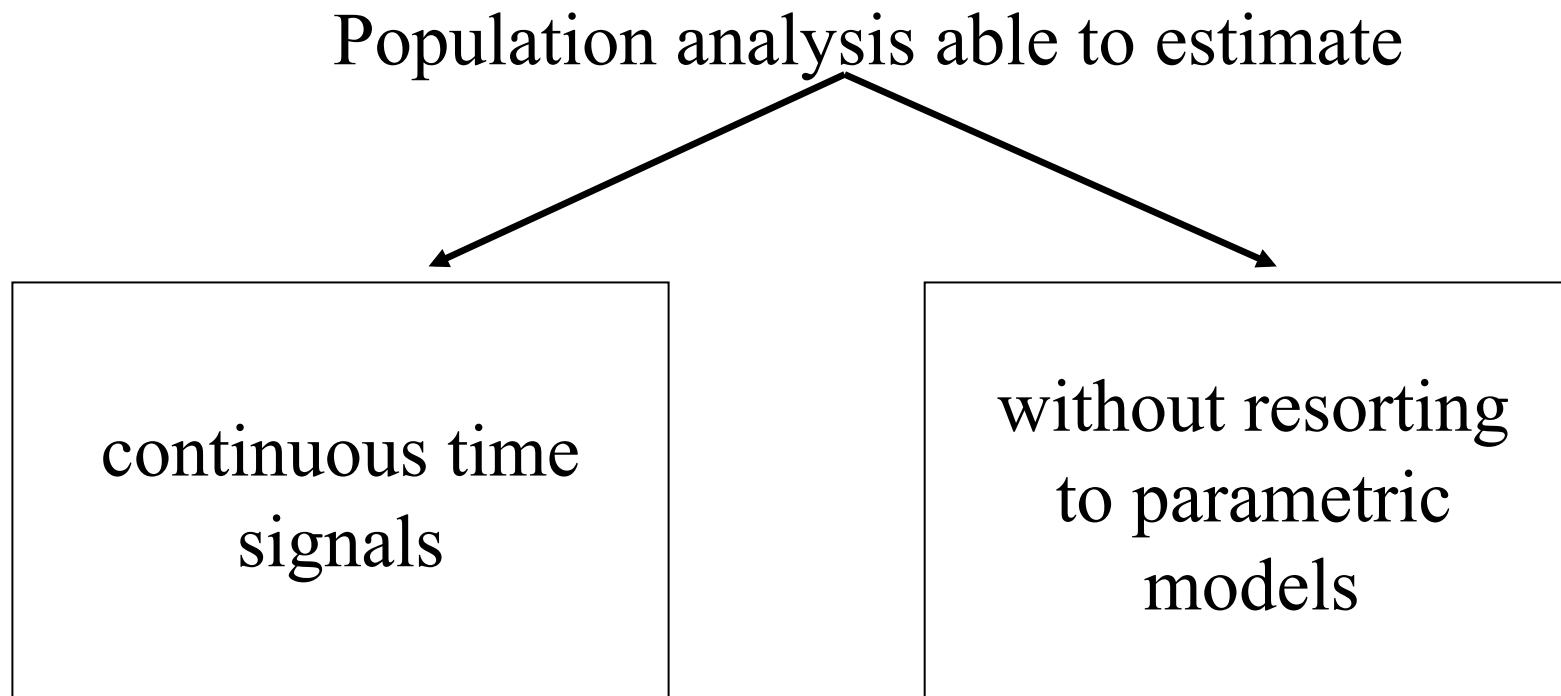
Drawback : number and location of spline knots

- ***Nonparametric models:***

Example: Random Walk [*Magni et al., 2002*]

Drawback : Discrete-time model

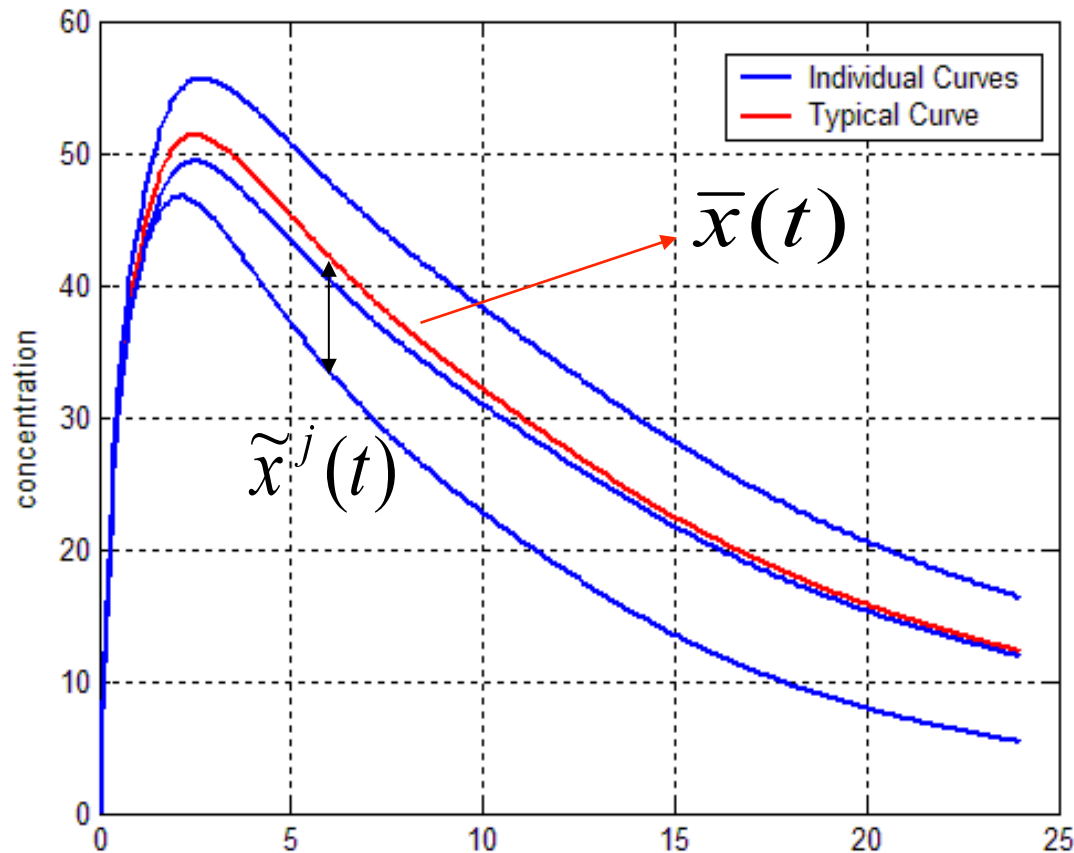
Open problems



Approach: Signals modelled as continuous-time stochastic processes

Parametric vs nonparametric

Population curves



$$x^j(t) = \bar{x}(t) + \tilde{x}^j(t)$$

The j -th individual curve $x^j(t)$ is given by the sum of:

- the typical curve:
 $\bar{x}(t)$

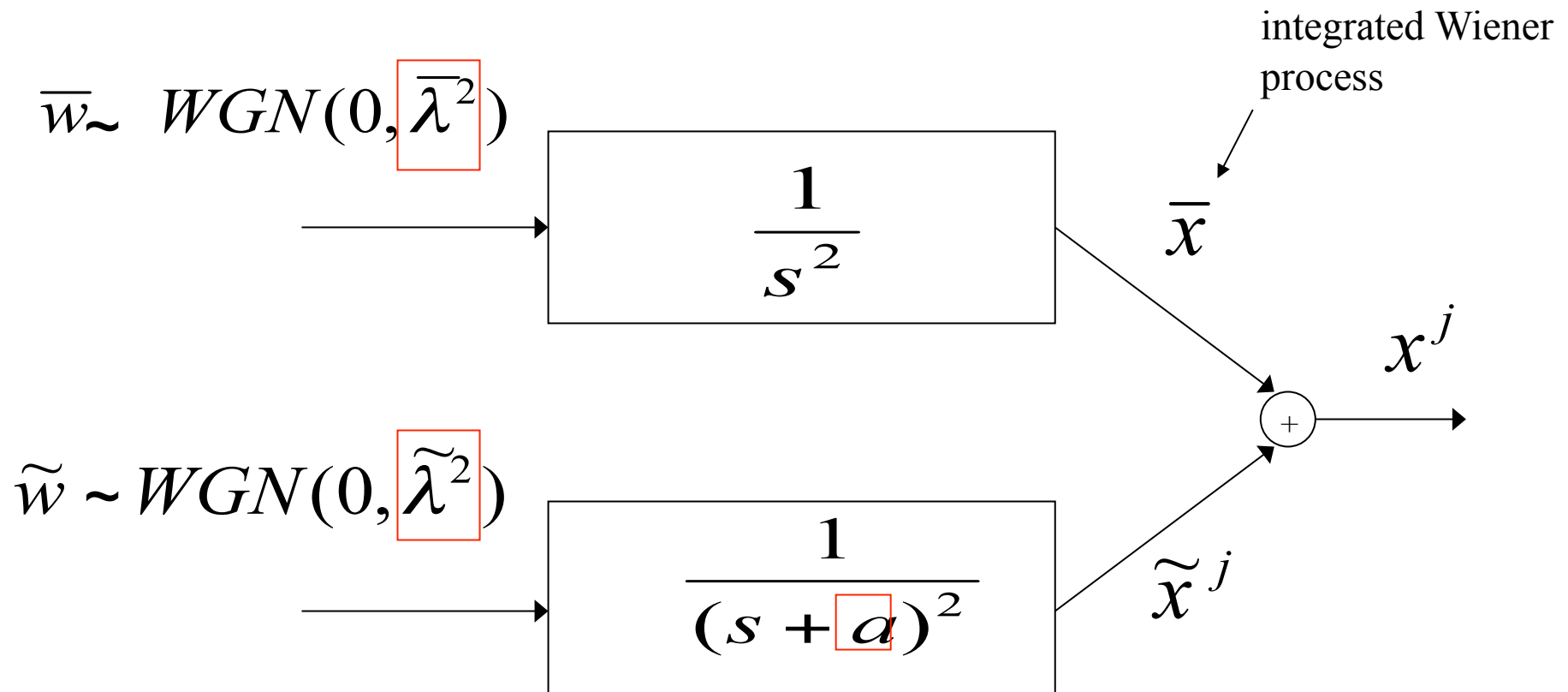
- the individual shift:

$$\tilde{x}^j(t)$$

Nonparametric model

Models and hyperparameters

Unknown hyperparameters:

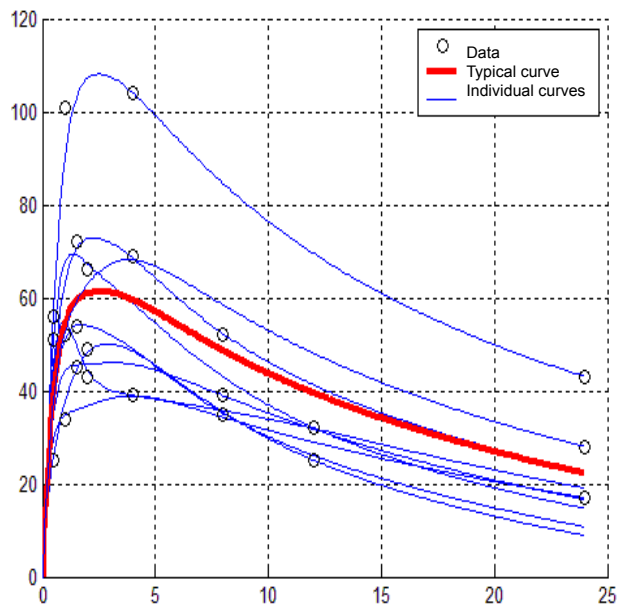


Assumption: zero initial state (can be relaxed)

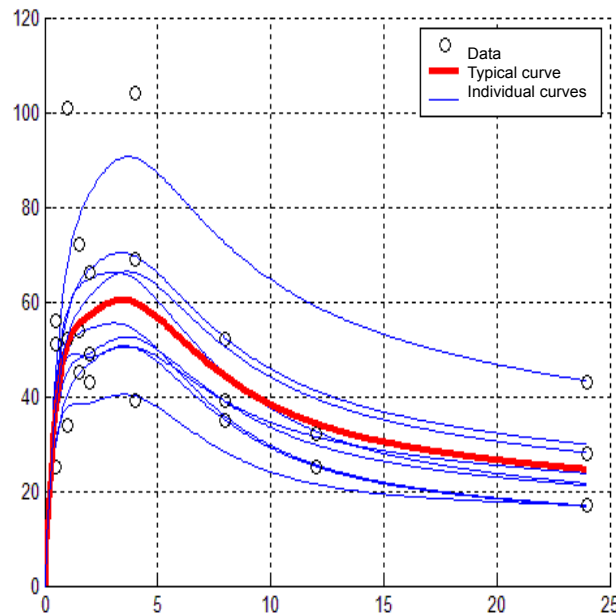
Nonparametric model

$\tilde{\lambda}^2$ effect

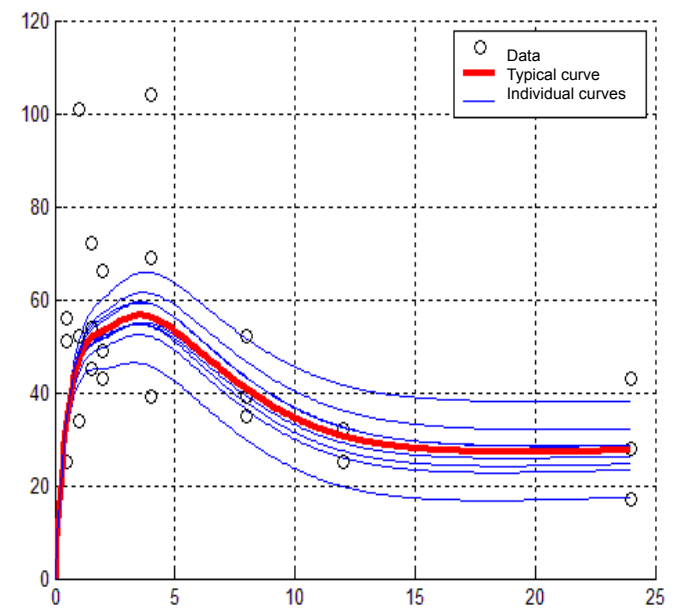
$$\tilde{\lambda}^2 = 10000$$



$$\tilde{\lambda}^2 = 1000$$



$$\tilde{\lambda}^2 = 100$$



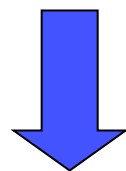
Nonparametric model

Signal Estimation

Bayesian estimation

Remarks :

- $\bar{x}(t)$ is a gaussian random variable with zero mean
- $Y = \bar{X} + \tilde{X} + V$ array of all the available data
- $(\bar{x}(t), Y)$ are jointly gaussian



$$\hat{\bar{x}}(t) = E[\bar{x}(t) | Y] = Cov[\bar{x}(t), Y] Var[Y]^{-1} Y$$

estimate of $\bar{x}(t)$

Signal estimation

Bayesian estimation

$$\bar{R}(t, \tau) = \bar{\lambda}^2 \bar{\gamma}(t, \tau) : \text{autocovariance of } \bar{x}(t)$$

$$\tilde{R}(t, \tau) = \tilde{\lambda}^2 \tilde{\gamma}(t, \tau) : \text{autocovariance of } \tilde{x}(t)$$

Remark 1:

$$\text{Cov}[\bar{x}(t), Y] = [\bar{R}(t, t_1) \dots \bar{R}(t, t_n)]$$

$$t_i \text{ (knots), } i = 1, \dots, n \quad n = n^\circ \text{ observations}$$

Remark 2:

$$\text{Var}[Y] \quad \text{depends on both } \bar{R}(t, t_i) \text{ and } \tilde{R}(t, t_i)$$

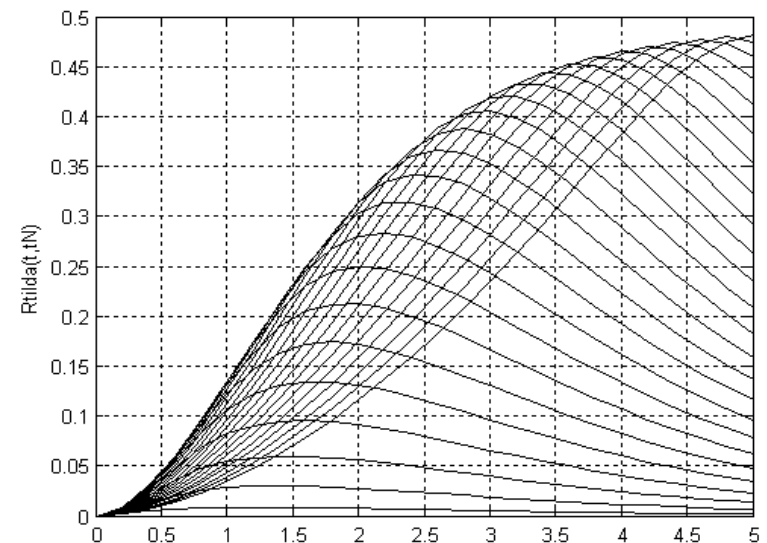
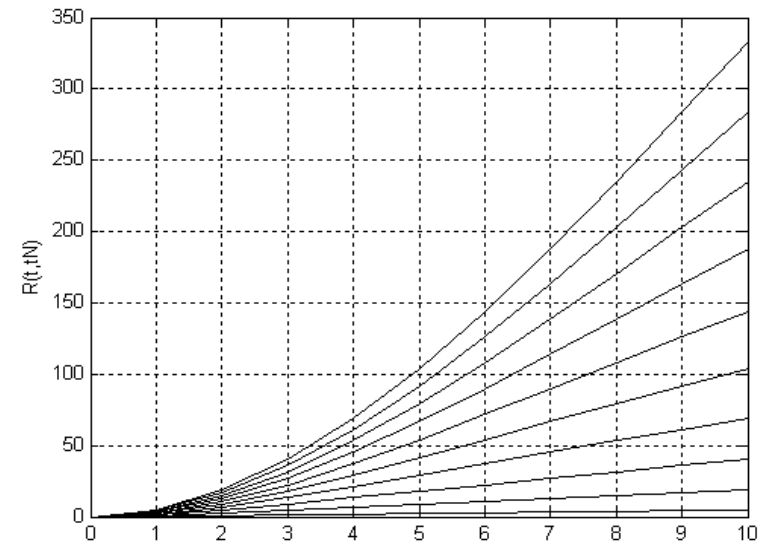
Autocovariance functions

Typical curve:

$$\bar{\gamma}(t, \tau) = \begin{cases} \frac{t^2}{2} \left(\tau - \frac{t}{3} \right) & t \leq \tau \\ \frac{\tau^2}{2} \left(t - \frac{\tau}{3} \right) & t > \tau \end{cases}$$

Individual shifts:

$$\tilde{\gamma}(t, \tau) = \begin{cases} e^{a(\tau-t)} \left(\frac{t^2 e^{2at}}{2a} - \frac{t \cdot e^{2at}}{2a^2} - \frac{e^{2at} - 1}{4a^3} \right) + \\ + (\tau - t) e^{a(\tau-t)} \frac{2at \cdot e^{2at} - e^{2at} + 1}{4a^2} & t \leq \tau \\ e^{a(t-\tau)} \left(\frac{\tau^2 e^{2a\tau}}{2a} - \frac{\tau \cdot e^{2a\tau}}{2a^2} - \frac{e^{2a\tau} - 1}{4a^3} \right) + \\ + (t - \tau) e^{a(t-\tau)} \frac{2a\tau \cdot e^{2a\tau} - e^{2a\tau} + 1}{4a^2} & t > \tau \end{cases}$$



Signal estimation

Estimate expressions

Hypothesis: known hyperparameters

*Estimate of the
typical curve:*

$$\hat{\bar{x}}(t) = \sum_{i=1}^n \bar{R}(t, t_i) \cdot c_i$$

→ Cubic
spline

*Estimate of the
individual
curve:*

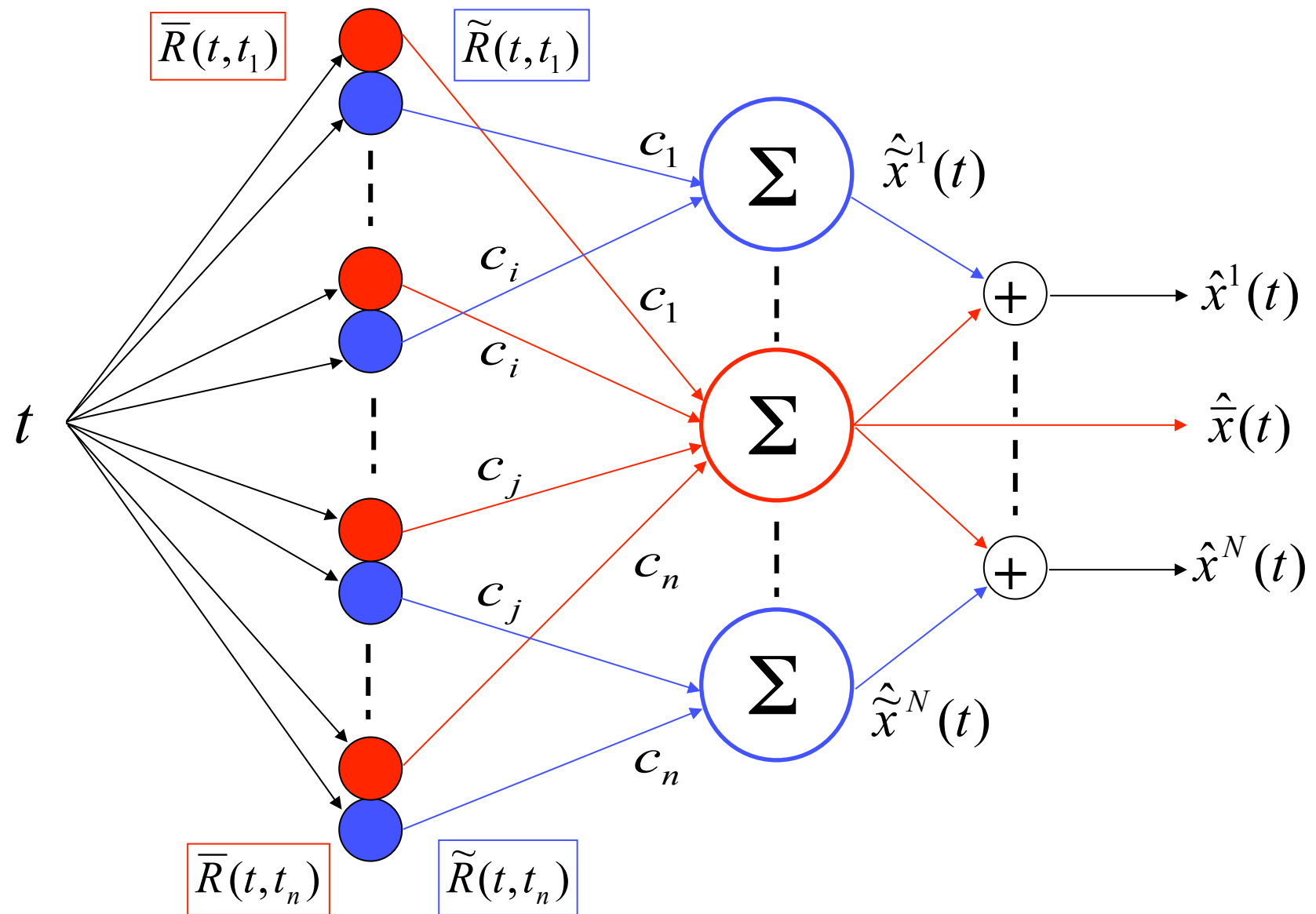
$$\hat{x}^j(t) = \sum_{i=1}^n \bar{R}(t, t_i) \cdot c_i + \sum_{i=1}^n \tilde{R}(t, t_i) \cdot c_i$$

Array of weights:

$$c = Var[Y]^{-1} Y$$

Signal estimation

Regularization Network



Signal estimation

*Hyperparameter Estimation:
Empirical Bayes and MCMC*

Empirical Bayes

Combination between

Classic approach:

Hyperparameter
estimation:
likelihood
maximization

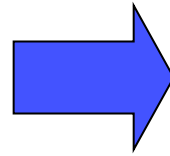
$$\theta = [\bar{\lambda}^2, \tilde{\lambda}^2, a]$$

$$\theta_{ML} = \arg \max(f(Y | \theta))$$

Bayesian approach:

Estimation of the
typical curve:

$$\hat{\bar{x}}(t) = E[\bar{x}(t) | Y, \theta = \theta_{ML}]$$



Hyperparameter estimation

MCMC algorithms

Bayesian Inference:

$$E[f(\theta) | Y] = \int_{\Theta} f(\theta) P(\theta | Y) d\theta$$

Problem: the integral does not admit an analytic solution



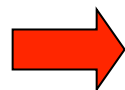
MONTE CARLO INTEGRATION

Hyperparameter estimation

MCMC algorithms

MONTE CARLO INTEGRATION

Given a sufficient
amount of samples
 $[\theta_1 \dots \theta_K]$ drawn from
the posterior:



$$E[f(\theta) | Y] \approx \frac{1}{K} \sum_{k=1}^K f(\theta_k)$$

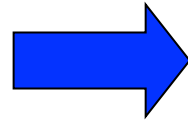
Problem: how to draw the samples?

Hyperparameter estimation

MCMC algorithms

MARKOV CHAINS

Construct a Markov chain which converges to the posterior distribution that has to be sampled



METROPOLIS-HASTINGS
(Gibbs Sampler)
algorithm

Hyperparameter estimation

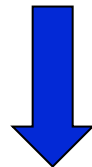
MCMC

Problem: MCMC simulation of the curves
in each time instant t



Computationally burdensome

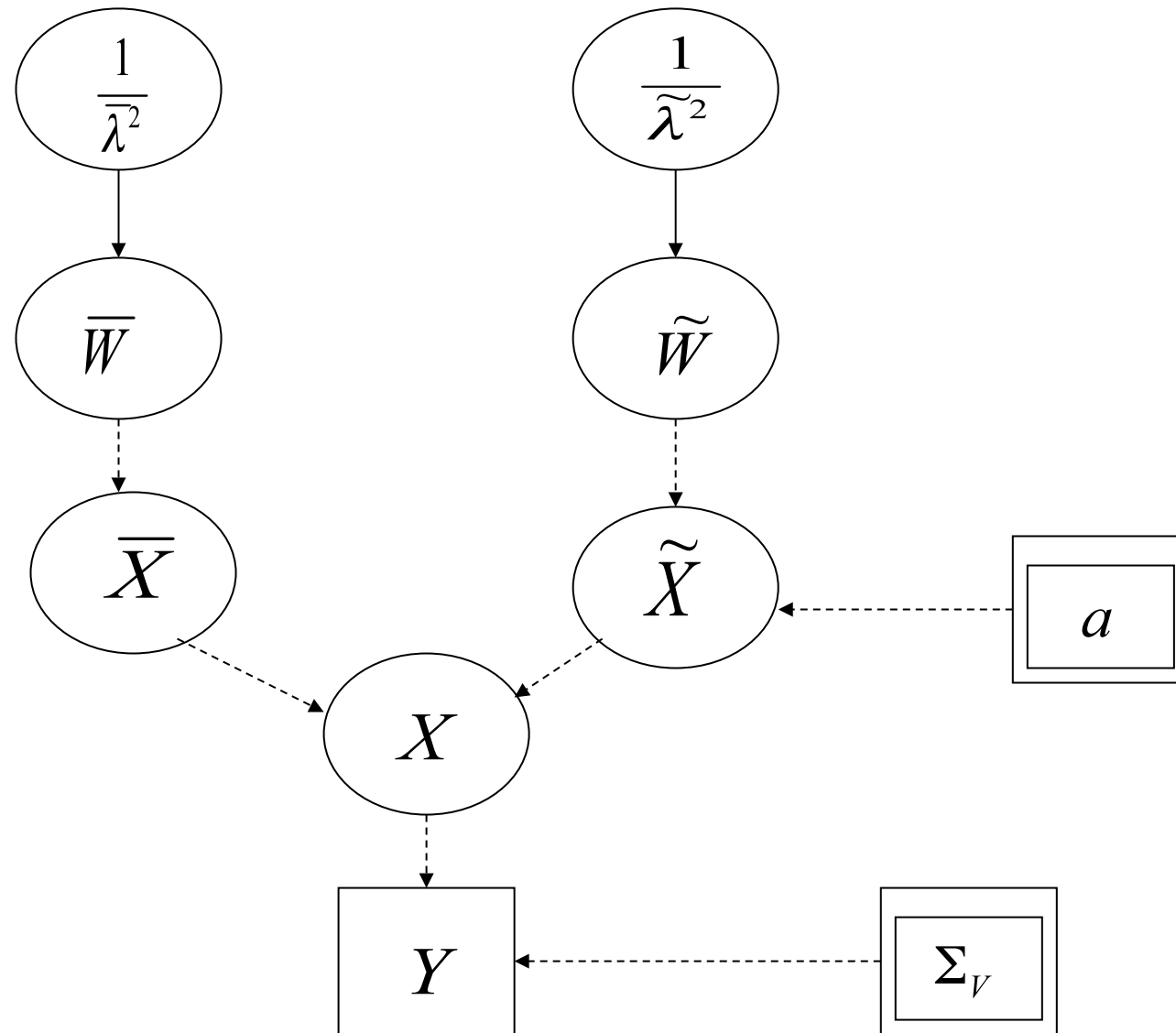
Solution: MCMC simulation of the curves
in correspondence with the nodes



only n random variables are sampled

Hyperparameter estimation

Bayesian network



Hyperparameter estimation

MCMC

Prior distributions:



$$p\left(\frac{1}{\bar{\lambda}^2}\right) = \Gamma(\gamma_1, \gamma_2)$$

$$p\left(\frac{1}{\tilde{\lambda}^2}\right) = \Gamma(\gamma_3, \gamma_4)$$

$$p(\bar{W}) = N(0, \bar{\lambda}^2 I_n)$$

$$p(\tilde{W}) = N(0, \tilde{\lambda}^2 I_n)$$

Full conditional:



$$p\left(\frac{1}{\bar{\lambda}^2} \mid \cdot\right) \propto \Gamma\left(\gamma_1 + \frac{n}{2}, \gamma_2 + \frac{\bar{W}^T \bar{W}}{2}\right)$$

$$p\left(\frac{1}{\tilde{\lambda}^2} \mid \cdot\right) \propto \Gamma\left(\gamma_3 + \frac{n}{2}, \gamma_4 + \frac{\tilde{W}^T \tilde{W}}{2}\right)$$

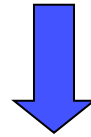
$$p(\bar{W}) \propto e^{-\frac{1}{2}\left(\frac{\bar{W}^T \bar{W}}{\bar{\lambda}^2}\right)} e^{\left(-\frac{1}{2}(Y - \bar{X})^T (\text{Var}[\tilde{X}] + \Sigma_V)^{-1} (Y - \bar{X})\right)}$$

$$p(\tilde{W}) \propto e^{-\frac{1}{2}\left(\frac{\tilde{W}^T \tilde{W}}{\tilde{\lambda}^2}\right)} e^{\left(-\frac{1}{2}(Y - \tilde{X})^T (\text{Var}[\bar{X}] + \Sigma_V)^{-1} (Y - \tilde{X})\right)}$$

Hyperparameter estimation

MCMC: the typical curve

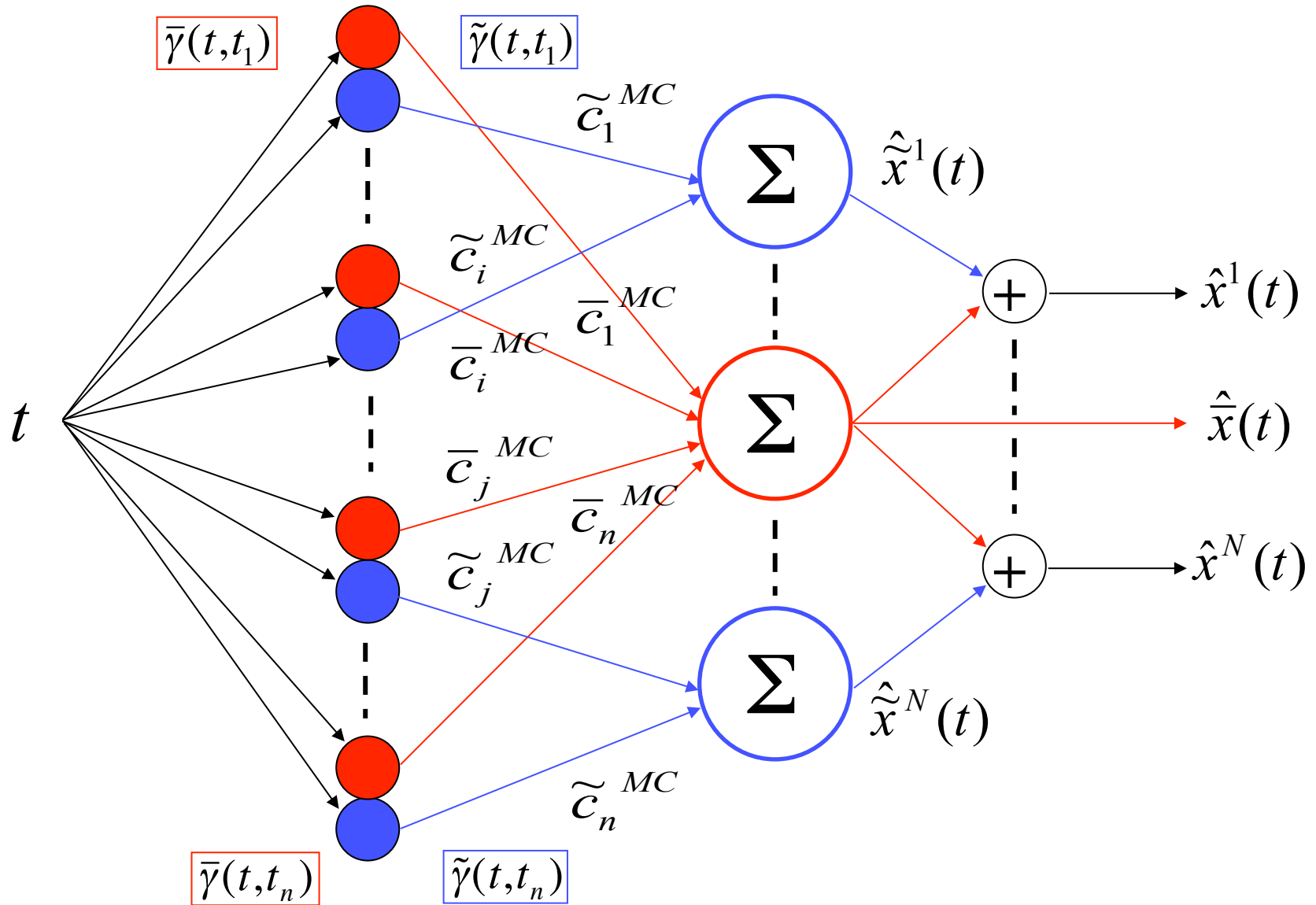
Given the posterior distribution: $p(\bar{\lambda}^2, \tilde{\lambda}^2 | Y)$
(hp: α known)



$$\begin{aligned}\hat{\bar{x}}(t) &= E[\bar{x}(t) | Y] = \int E[\bar{x}(t) | Y, \theta] p(\theta | Y) d\theta \\ &\cong \sum_{k=1}^K \frac{E[\bar{x}(t) | Y, \bar{\lambda}_k^2, \tilde{\lambda}_k^2]}{K} = \sum_{i=1}^n \bar{\gamma}(t, t_i) \cdot \bar{c}_i^{MC}\end{aligned}$$

with $\bar{c}_i^{MC} = \sum_{k=1}^K \frac{\bar{\lambda}_k^2 c_i(\bar{\lambda}_k^2, \tilde{\lambda}_k^2)}{K}$ (K samples)

Modified Regularization Network

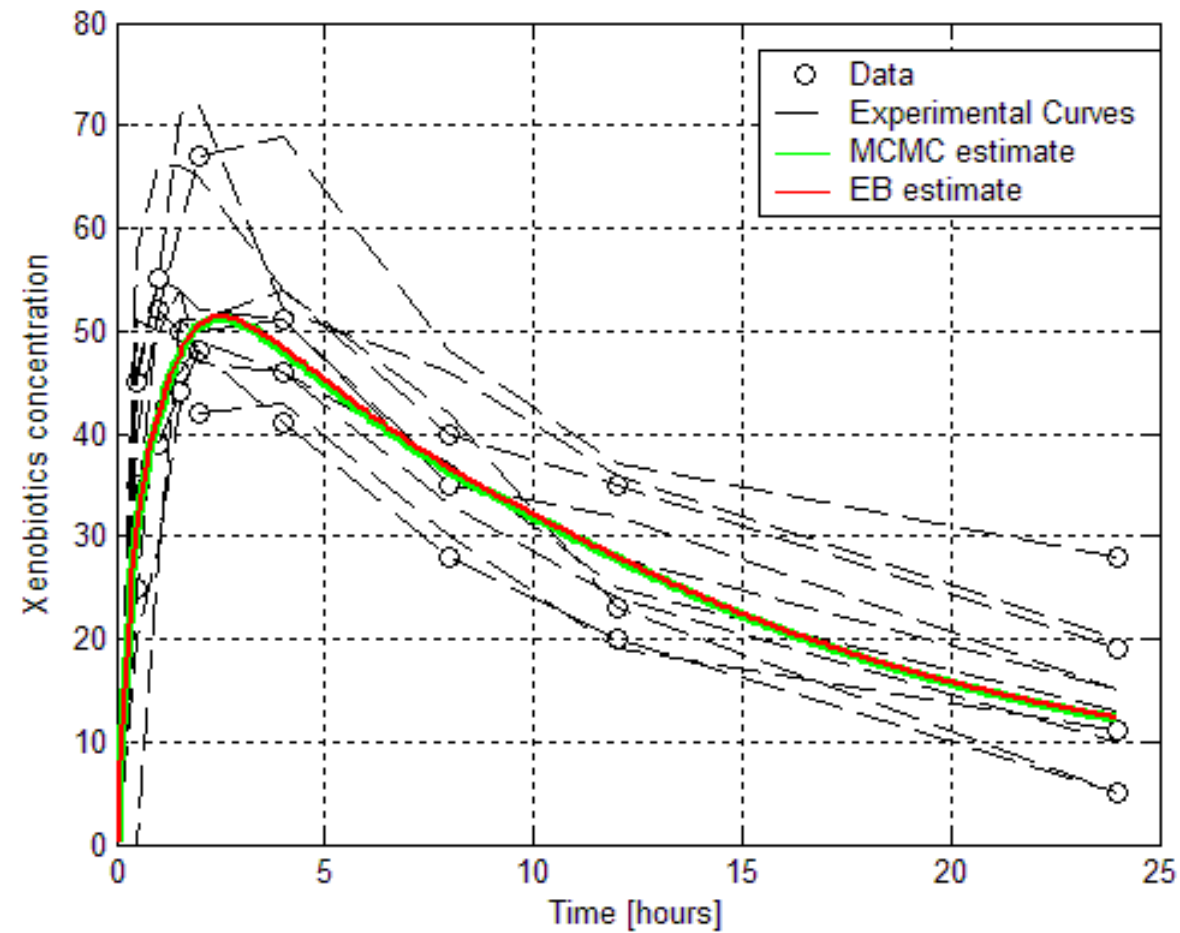


Hyperparameter estimation

Application to PK Data

Pharmacokinetic data

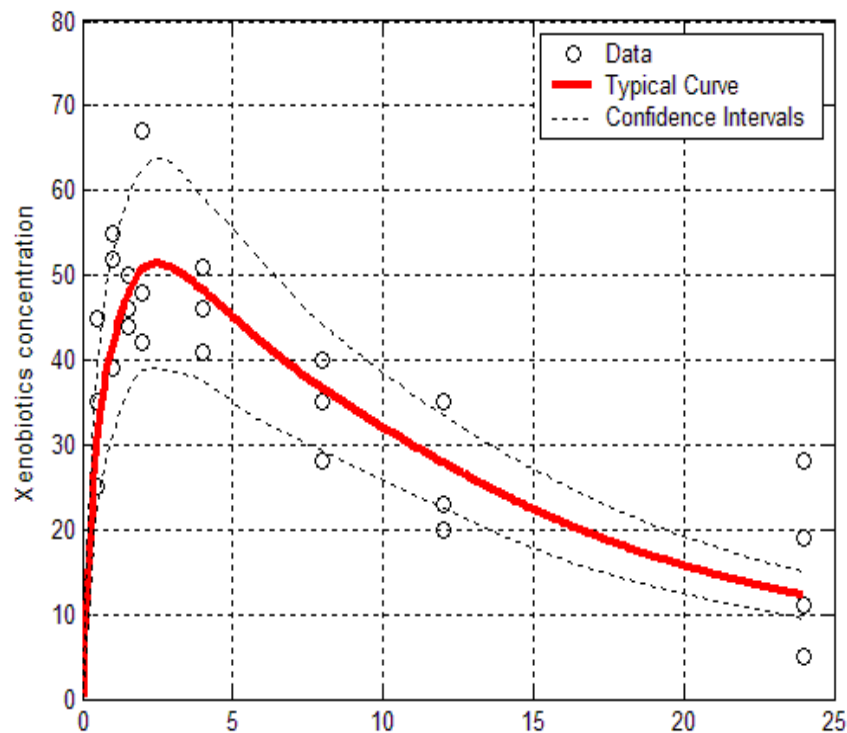
$N=10$ subjects



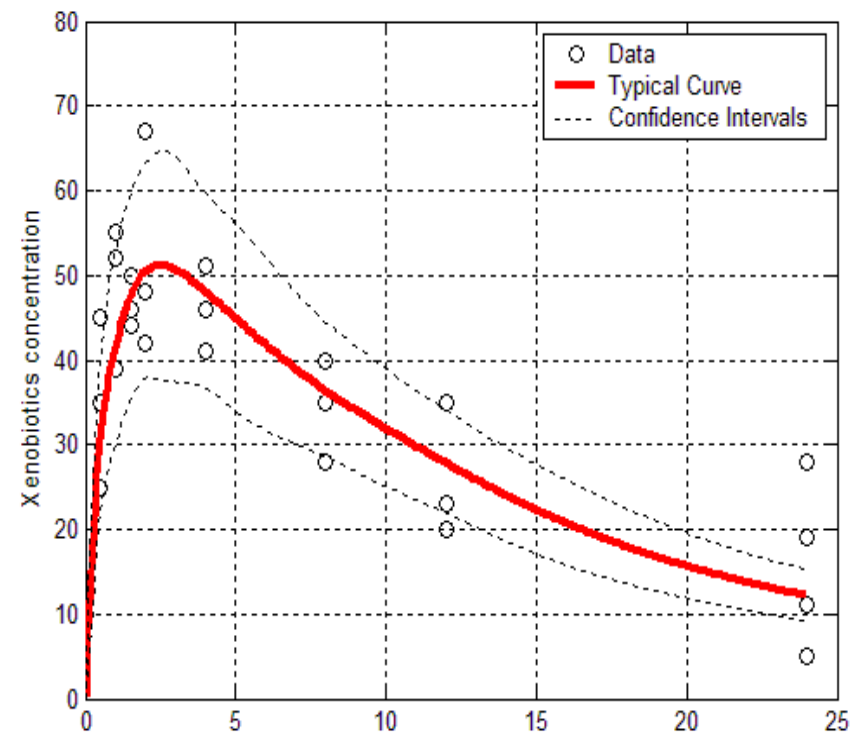
Application to PK data

Typical curve estimates

EB estimate



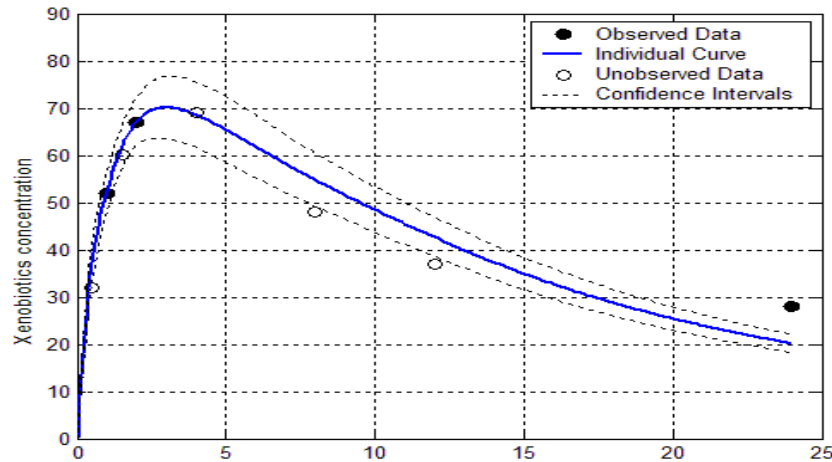
MCMC estimate



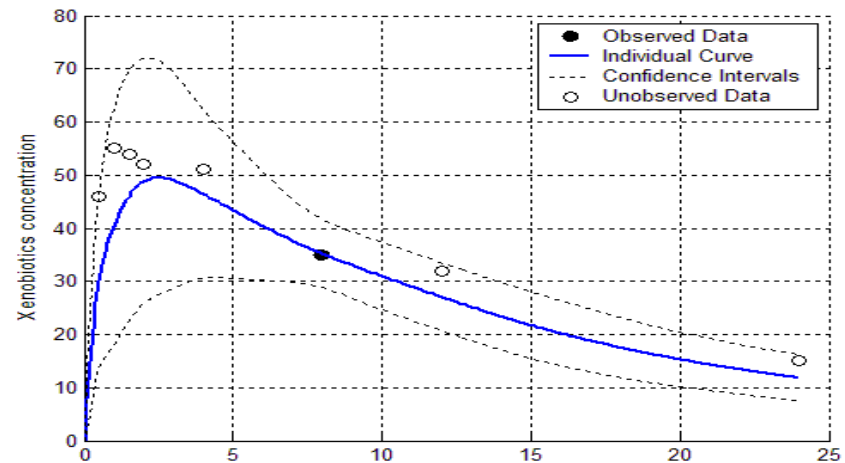
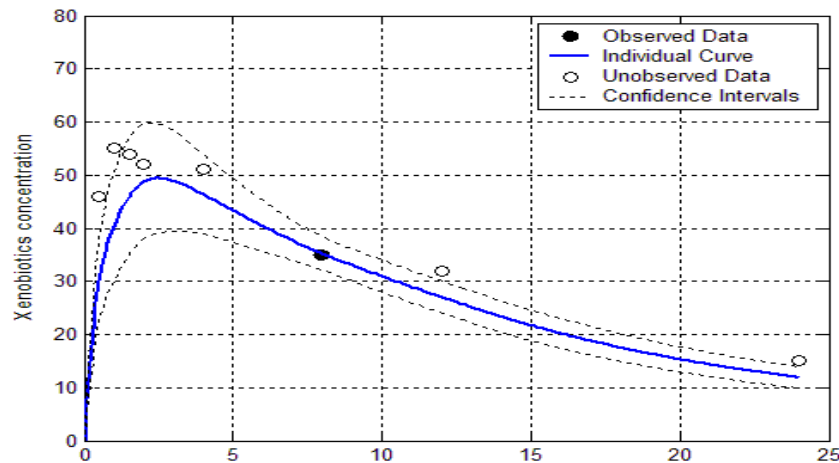
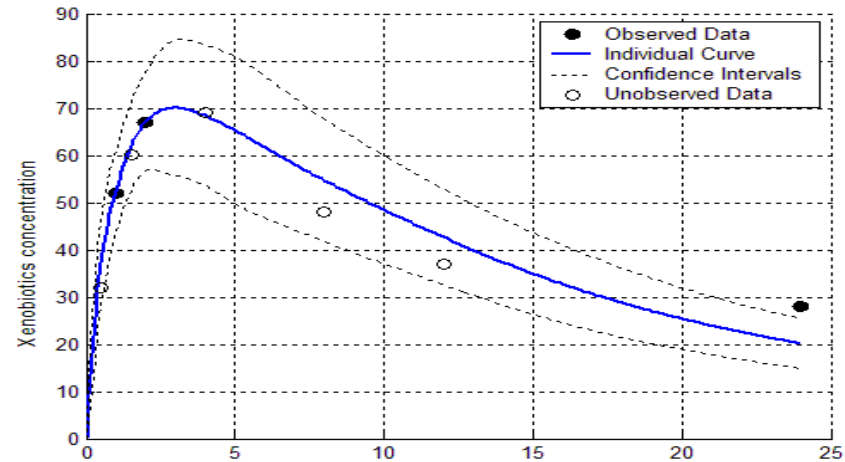
Application to PK data

Individual curve estimates

EB estimate



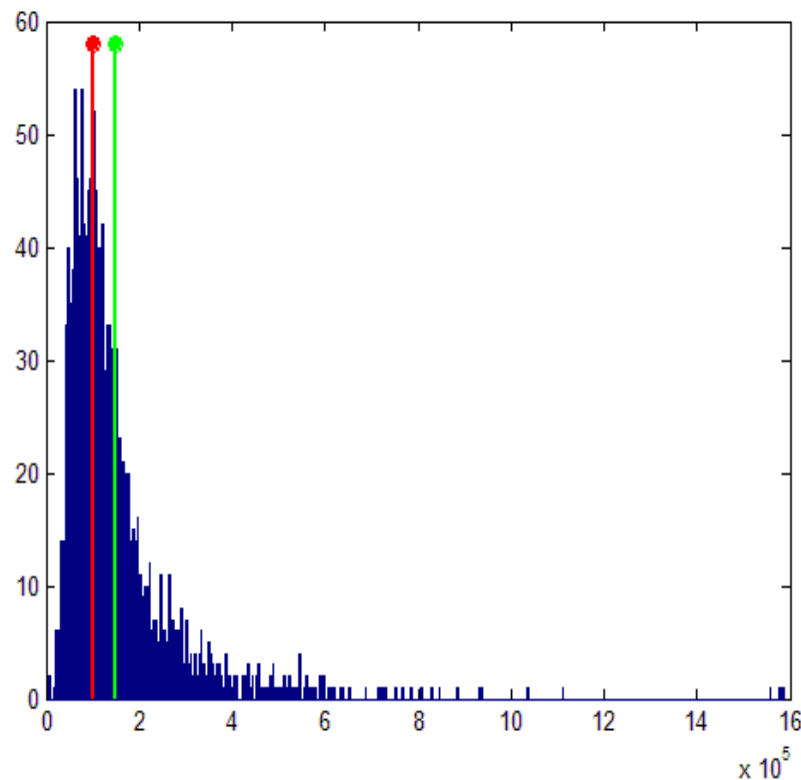
MCMC estimate



Application to PK data

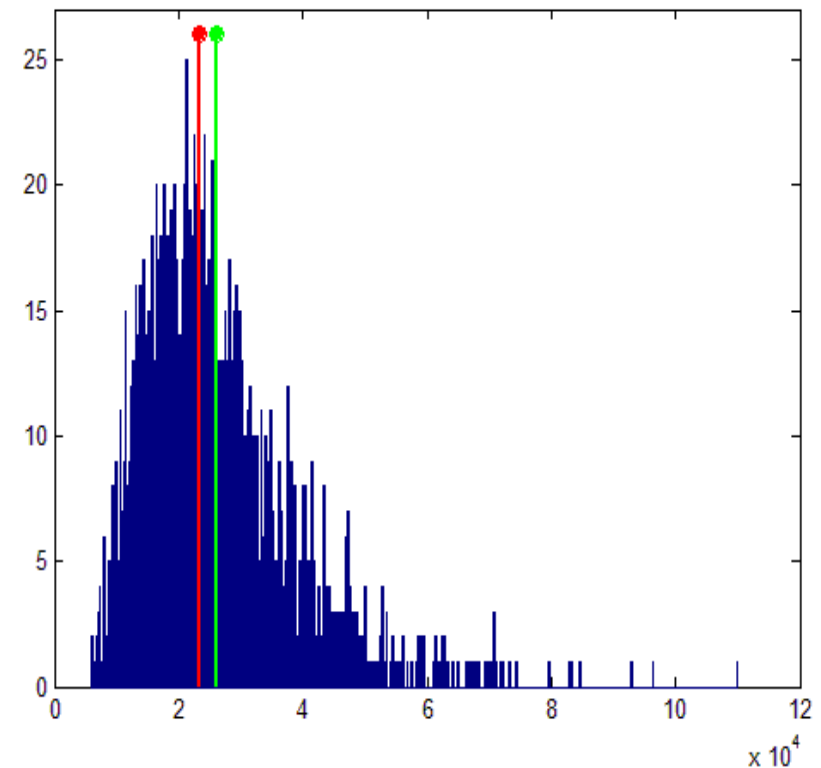
Hyperparameter posterior

Histogram of $\overline{\lambda}^2$



● EB estimate

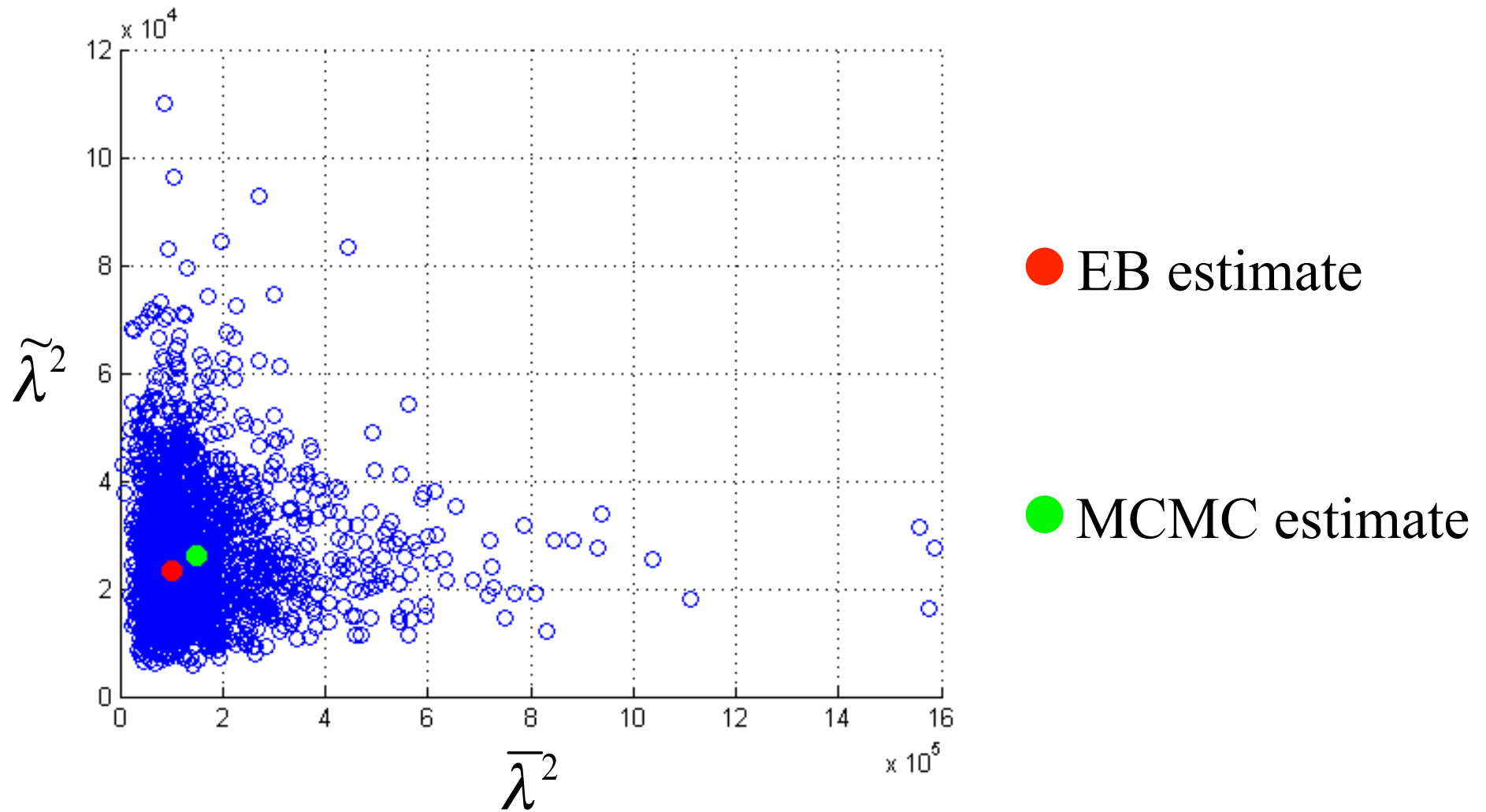
Histogram of $\tilde{\lambda}^2$



● MCMC estimate

Application to PK data

Scatter plot of the couples $(\bar{\lambda}^2, \tilde{\lambda}^2)$



Application to PK data

Conclusions

- Continuous- time estimates solving a finite dimensional problem
 - Estimate of the typical curve: cubic spline
 - Regularization Network structure
 - Hyperparameter estimation: Empirical Bayes and MCMC (the latter does not underestimate confidence intervals)
 - MCMC estimates for the intersample without drawing additional intermediate samples
-