



# AN OVERVIEW OF SUBSPACE IDENTIFICATION

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Stockholm - May 11th, 2007



# Outline of the Talk

- Identification
  - Statement of the problem
  - Closed loop identification
- Subspace identification and stochastic realization
  - Motivations
  - Main ideas
- Main (classical) Algorithms (CCA, N4SID, MOESP)
- Closed Loop Subspace Identification (New Alg.)
- Recent contributions
  - Statistical Analysis (Asymptotic Variance)
  - Comparison of algorithms
  - Role of the Future Horizon
- Simulation results
- Conclusions
  - Open questions



# What I Shall not discuss...

- Identification of cointegrated process  
*[Bauer-Wagner]*
- ARMA processes with conditionally heteroskedastic innovation  
(e.g. (G)ARCH processes) *[Bauer]*
- Order Estimation *[Sorelius, Peterzell et al., Bauer, Camba-Mendez et al.....]*



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# Identification of LTI systems

Given observed data  $\{y_1, \dots, y_T, u_1, \dots, u_T\}$ ,  $y_t \in \mathbb{R}^{p_y}$ ,  $u_t \in \mathbb{R}^{m_u}$ , construct (approximate) model

Model order will always be  $n$

$$\begin{cases} \mathbf{x}(t+1) &= \mathbf{A}\mathbf{x}(t) + \mathbf{B}\mathbf{u}(t) + \mathbf{K}\mathbf{e}(t) \\ \mathbf{y}(t) &= \mathbf{C}\mathbf{x}(t) + \mathbf{e}(t) \end{cases}$$

(LTI)

which “describes” the data ( $\mathbf{e}(t)$  is “white noise”)

If you prefer transfer functions

$$\mathbf{y}(t) = \mathbf{F}(z)\mathbf{u}(t) + \mathbf{G}(z)\mathbf{e}(t)$$

“Innovation” = one step ahead prediction error

where

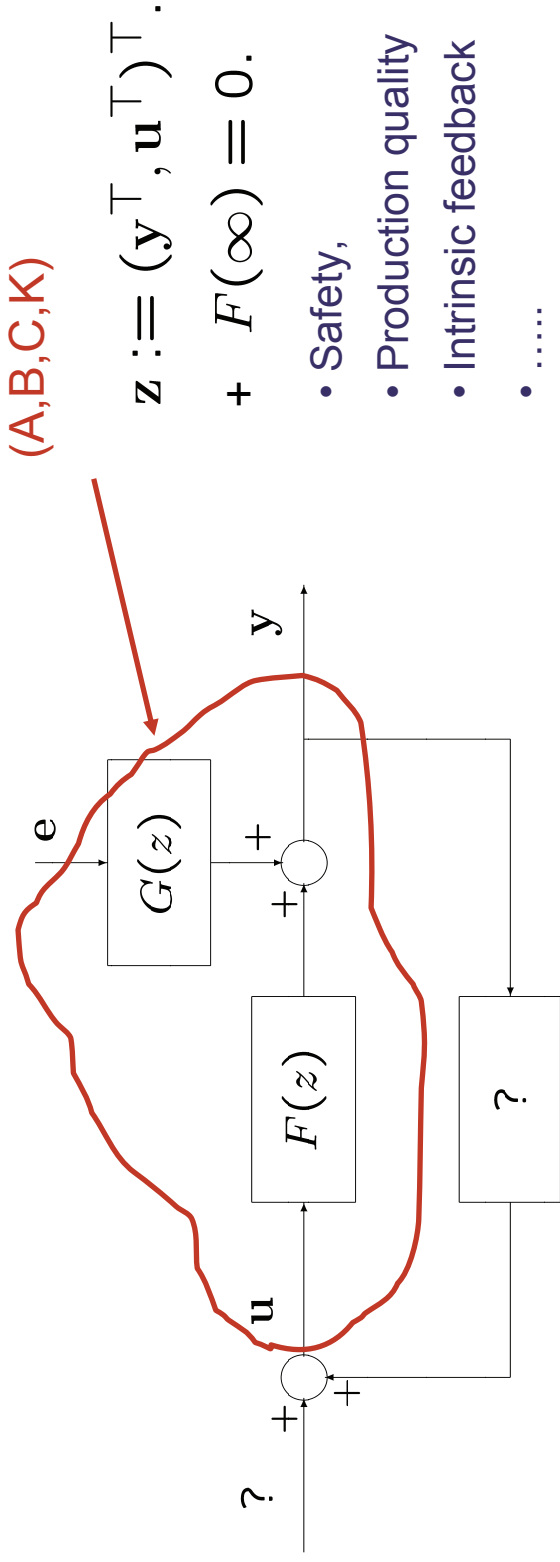
$$\mathbf{F}(z) := \mathbf{C}(z\mathbf{I} - \mathbf{A})^{-1}\mathbf{B} \quad \mathbf{G}(z) := \mathbf{C}(z\mathbf{I} - \mathbf{A})^{-1}\mathbf{K} + \mathbf{I}$$

“deterministic”

“stochastic”



# Identification (..closed loop)



$$\mathbf{z} := (\mathbf{y}^\top, \mathbf{u}^\top)^\top.$$

$$+ F(\infty) = 0.$$

- Safety,
- Production quality
- Intrinsic feedback
- .....

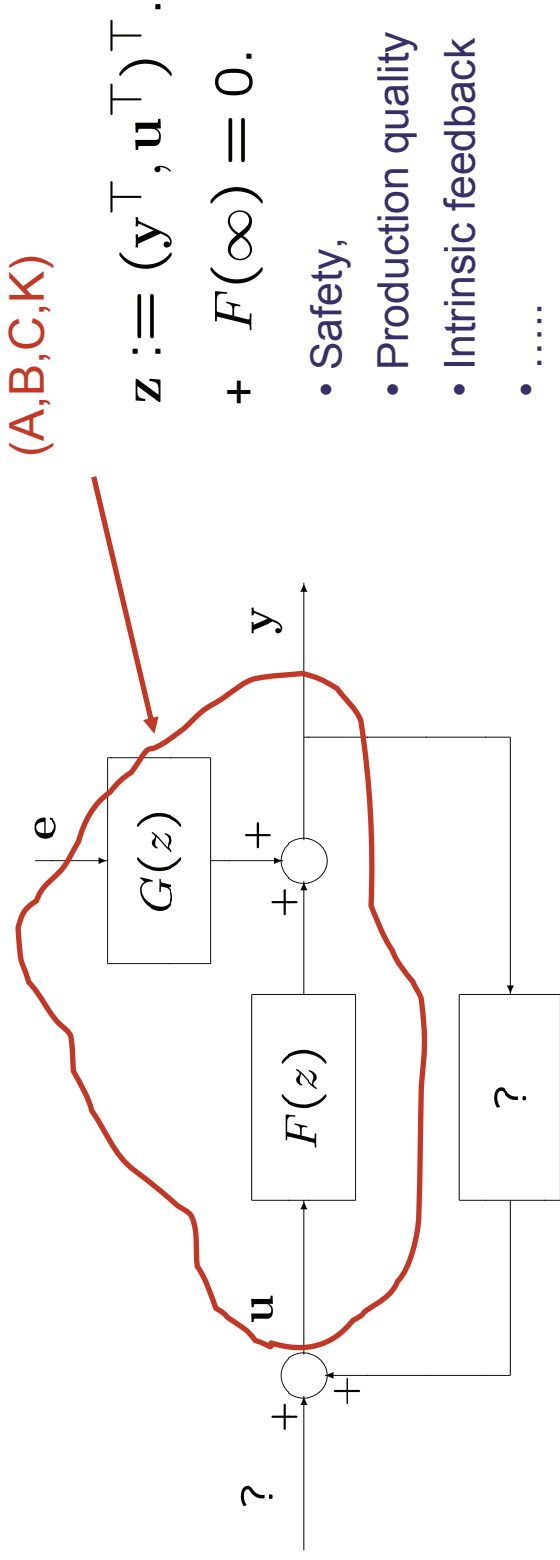
- Sometimes we shall need

$$mI \leq \Phi_{\mathbf{z}}(z) \leq MI, \quad 0 < m \leq M < \infty$$

Joint Spectrum



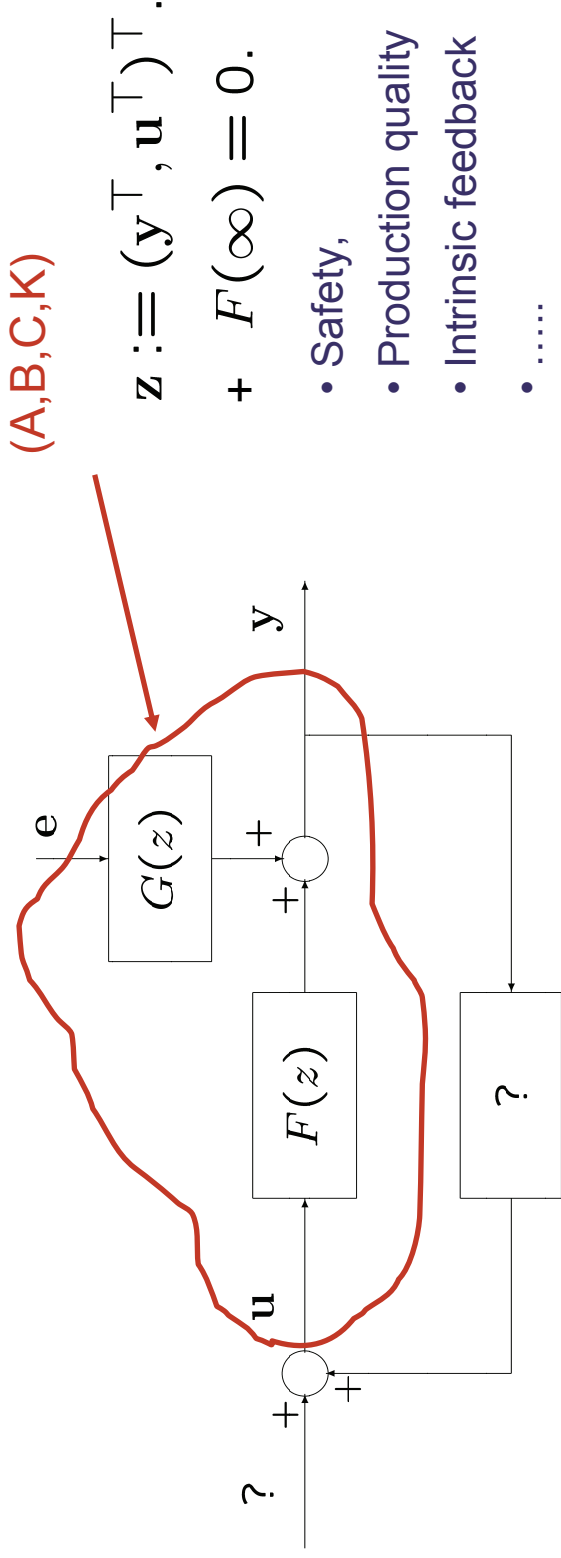
# Identification (..closed loop)



- Estimate Parameters (A,B,C,K) [Point estimators in statistics]
- Attach “quality tag” to the estimators [e.g. Variance]



# Identification (..closed loop)



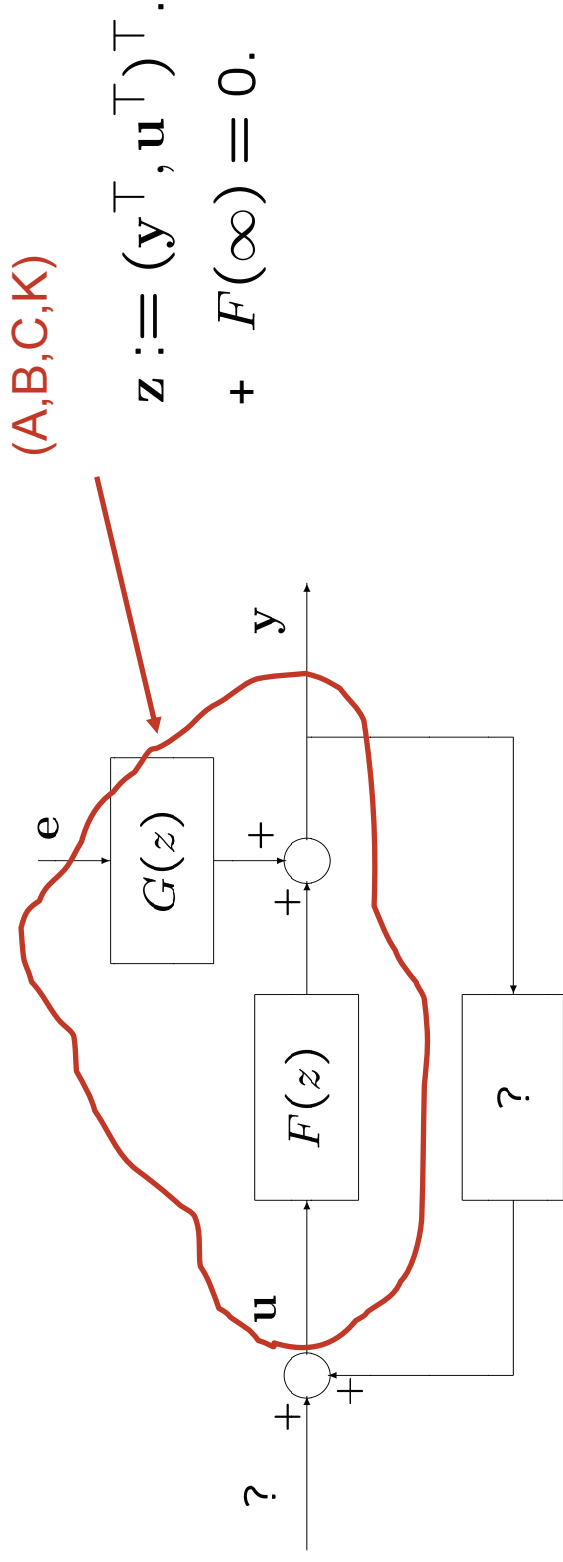
- Estimate Parameters (A,B,C,K) [Point estimators in statistics]
- Attach “quality tag” to the estimators [e.g. Variance]

WELL-POSED ???





# Identification (..closed loop)



“Essentially” 1 : 1 correspondence between the covariance sequence of the joint process  $(y, u)$  and internally stable (linear) feedback models  $\rightarrow$  UNIQUE  $F(z)$ !

[Ljung-Soderström-Gustavsson-Gevers-Anderson-Caines-Chan]  
[1974-1981]



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- **Subspace identification and stochastic realization**
  - Motivations (i.e. WHY?)
  - Main ideas (i.e. WHAT IS IT?)
- Main (classical) Algorithms (CCA,N4SID,MOESP)
- Closed Loop Subspace Identification (New Alg.)
- Recent contributions
  - Statistical Analysis (Asymptotic Variance)
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# WHY Subspace ID? (and why not!)

- **PROS:**
  1. No parametric optimization (as PEM), only linear algebra (QR, SVD,..) (+ Riccati)
  2. Implements Stochastic realization procedures (*understand what's going on!*)
  3. Identification with feedback (*Recent !*)  
→ MIMO identification (APPLICATIONS)
- **CONS:**
  1. Statistical Analysis Difficult
  2. Many algorithms, relative efficiency not completely clear + finite data behavior



# A 60x60 MIMO real system...

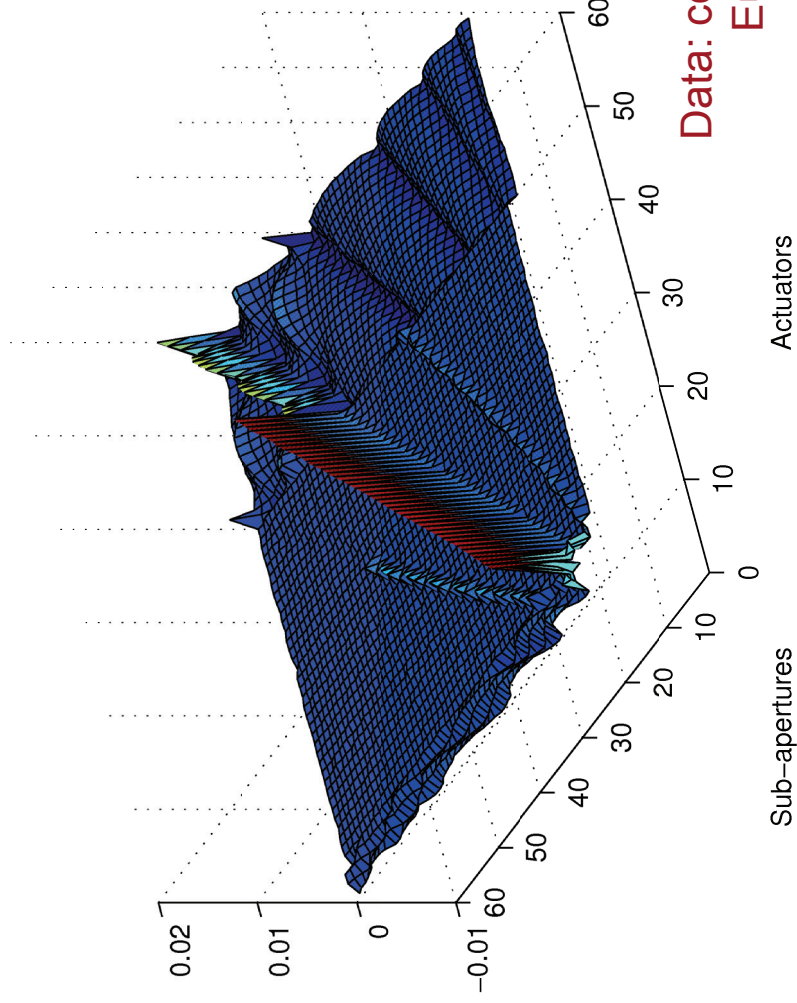
- Closed loop Identification of the Deformable Mirror in adaptive optics (DC Gain shown)



# A 60x60 MIMO real system...

- Closed loop Identification of the Deformable Mirror in adaptive optics (DC Gain shown)

Interaction Matrix (DC gain of identified model)



Identification of the “full”  
60x60 system with PBSIDopt  
Using closed loop data  
(i.e. when the mirror curvature is  
controlled)

Matches very well with the  
DC gain which can be computed  
from a very accurate  
(non-linear) model of the plant and  
with physical insight

Data: courtesy of

European Southern Observatory (ESO)

Sub-apertures

Actuators

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# Models from covariances

Note: given  $\{z_t\}_{t \in \mathbb{R}}$ ,  $z_t = [y_t^\top u_t^\top]^\top \rightarrow \Sigma(\tau)$  (Ergodicity)

**Problem:** Given  $\Sigma(\tau)$  find  $A, B, C, K, \Lambda = \text{Var}\{e(t)\}$   
compute a (minimal) state space model (LTI).



## STOCHASTIC REALIZATION PROBLEM

Huge liter.: *Akaike, Ruckebusch, Faurre, Lindquist and Picci, Chiuso and Picci*  
(Just to cite a few!)

**No** measured “inputs”  $u(t)$   
(time-series identification)  
(1969-1974-1985-1991)

Measured “inputs”  $u(t)$   
(identification with inputs)  
(2002-2005)



# Geometric Stochastic Realization

Given a stat. stoch. process  $\{\mathbf{z}(t)\}_{t \in \mathbb{R}}$ ,  $\mathbf{z}(t) = [\mathbf{y}^\top(t), \mathbf{u}^\top(t)]^\top$

Let  $\mathcal{H} := \overline{\{\eta := \sum_i a_i^\top \mathbf{z}(t_i)\}} = \overline{\text{span}\{\mathbf{z}(s), s \in \mathbb{Z}\}}$



# Geometric Stochastic Realization

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Similar considerations in  
Willems (1986-2006) in a  
deterministic setting

## Geometric Stoch. Realization:

1. construct  $\mathcal{X}_t := \text{span}\{\mathbf{x}(t)\}$  via “geometric” operations on the space  $\mathcal{H}$  (COVARIANCE SEQUENCE).
2. to each (many !!!) “state space” one can attach a realization  $(A, B, C, D)$  choosing a basis  $\mathbf{x}(t)$  (equivalence class as for deterministic realization)





# From Stochastic Realization to Subspace Algorithms

$$y(t) \xleftrightarrow[N \rightarrow \infty]{1:1} \{y_t, y_{t+1}, \dots, y_{t+N-1}\} = \mathbf{Y}_t$$

*Stoch. Processes*  $\leftrightarrow$  *Lin. Algebra*



# From Stochastic Realization to Subspace Algorithms

$$y(t) \xrightleftharpoons[N \rightarrow \infty]{1:1} \{y_t, y_{t+1}, \dots, y_{t+N-1}\} = Y_t$$

*Stoch. Processes*  $\leftrightarrow$

*Lin. Algebra*

QR, SVD...

$$\mathcal{Y}_{[t,s]} \stackrel{||}{=} \text{span}\{y(t), \dots, y(s)\} \xrightleftharpoons[N \rightarrow \infty]{1:1} Y_{[t,s]} :=$$

$$\begin{bmatrix} y_t & y_{t+1} & \dots & y_{t+N-1} \\ y_{t+1} & y_{t+2} & \dots & y_{t+N} \\ \vdots & \vdots & \ddots & \vdots \\ y_s & y_{s+1} & \dots & y_{s+N-1} \end{bmatrix}$$

(abstract) Spaces of random variables  $\leftrightarrow$  Row space of (Block) Hankel Data Matrices



# Key Steps in Subspace ID

→ sometimes called past induced.....

1. Construct a basis for the (predictor) state space  $\mathbf{X}_t$   
(a matrix!!)

2.  $\hat{A}$ ,  $\hat{B}$ ,  $\hat{C}$ ,  $\hat{K}$  (+ Riccati...) FROM

$$\begin{cases} \mathbf{X}_{t+1} = A\mathbf{X}_t + B\mathbf{U}_t + K\hat{\mathbf{E}}_t \\ \mathbf{Y}_t = C\mathbf{X}_t + \mathbf{E}_t \end{cases}$$

where  $\hat{\mathbf{E}}_t := \mathbf{Y}_t - \hat{C}\mathbf{X}_t$        $\frac{\mathbf{E}_t\mathbf{X}_t^\top}{N} \simeq \mathbb{E}[\mathbf{e}(t)\mathbf{x}^\top(t)] = 0$



# How do we get the state?

1. The state  $\mathcal{X}_p$  is the “interface” between “past”  $(-\infty, p)$  and “future”  $[p, +\infty)$  (**=SPLITTING**)
2. (*no  $\mathbf{u}(t)$* ) Every future output can be decomposed in a part that lies in the past (and hence in the state) and one that is “uncorrelated” with it (minimum variance estimator)

$$y(p+k) = \hat{y}(p+k|p) + \tilde{y}(p+k|p)^\perp$$

where

$$\hat{y}(p+k|p) := E[y(p+k)|\mathcal{Y}_p^-] = CA^k \mathbf{x}(p) \in \mathcal{X}_p$$

Orthogonal projection  
(minimum variance predictor)

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# The state space (no inputs)

1. A little of work will show that

$$\mathcal{X}_p := E [\mathcal{Y}_p^+ | \mathcal{Y}_p^-]$$

Orthogonal projection

is a “valid state space” (**Markovian Splitting**)

2. How do we construct it in practice?

Canonical correlation analysis (CCA) comes into play: allows to discover **subsets** of variables in  $\mathcal{Y}_p^-$  that are “maximally correlated” with  $\mathcal{Y}_p^+$ .  
(**model reduction/approximation..**)

Hotelling, 1936: *Relations between two sets of variables*



# CCA (no inputs)

1. Compute

p=length of ``past``  
f+1=length of ``future``

$$\hat{\Sigma}_{fp} := \frac{1}{N} \mathbf{Y}_{[p,p+f]} \mathbf{Y}_{[0,p]}^T$$

$$\hat{\Sigma}_{pp} := \frac{1}{N} \mathbf{Y}_{[0,p]} \mathbf{Y}_{[0,p]}^T$$

$$\hat{\Sigma}_{ff} := \frac{1}{N} \mathbf{Y}_{[p,p+f]} \mathbf{Y}_{[p,p+f]}^T$$



# CCA (no inputs)

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  - $p$ =length of ``past``
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2. Then SVD decompose

$$U_n \mathbf{S}_n \mathbf{V}_n^T \simeq \hat{\Sigma}_{ff}^{-1/2} \hat{\Sigma}_{fp} \hat{\Sigma}_{pp}^{-1/2}$$



# CCA (no inputs)

p=length of ``past``  
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1. Compute

$$\begin{aligned} \hat{\Sigma}_{fp} &:= \frac{1}{N} \mathbf{Y}_{[p,p+f]} \mathbf{Y}_{[0,p]}^T \\ \hat{\Sigma}_{pp} &:= \frac{1}{N} \mathbf{Y}_{[0,p]} \mathbf{Y}_{[0,p]}^T \\ \hat{\Sigma}_{ff} &:= \frac{1}{N} \mathbf{Y}_{[p,p+f]} \mathbf{Y}_{[p,p+f]}^T \end{aligned}$$

2. Then SVD decompose

$$U_n S_n V_n^T \simeq \hat{\Sigma}_{ff}^{-1/2} \hat{\Sigma}_{fp} \hat{\Sigma}_{pp}^{-1/2}$$

3. Set

$$\hat{\mathbf{X}}_p := S_n^{-1/2} U_n^T \hat{\Sigma}_{ff}^{-1/2} \hat{\Sigma}_{fp} \hat{\Sigma}_{pp}^{-1} \mathbf{Y}_{[0,p]}$$





# With Inputs/Feedback

- Similar ideas/more complicated formulas
- Involves “oblique projections” and conditional CCA
- Feedback? Needs new point of view:  
**the predictor model**  
(rather simple idea to be honest!)



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# Milestones in Subspace ID

- Akaike (1975-76) Canonical Variables as “state”.
- Moonen et al. (1988) Deterministic ID: “*Intersection algorithms*” (in the spirit of Willems 1986)
- Aoki (1990): Realization applied to estimated covariance (*a la Ho-Kalman*)
- Larimore (1983-1990): first “**CCA**” method
- Verhaegen et al. (1992-1994): **MOESP** “class” (work on observability matrix rather than state – “**shift invariance**”)
- van Overschee - De Moor (1993): first time “finite-interval stochastic realization translated right!” (time-series id.) [**CCA + Riccati to compute KJ**]
- van Overschee - De Moor (1994): **N4SID** algorithm
- Picci-Katayama, Chiuso-Picci (1996-1999) – **ORT DEC Algorithm**

Box-Jenkins sort of model



# NOTATION

Observability matrix

$$\Gamma := \begin{bmatrix} C \\ CA \\ CA^2 \\ \vdots \\ CA^f \end{bmatrix}$$

“deterministic” ( $u$  as input)

$$H_d = \begin{bmatrix} 0 & 0 & 0 & \dots & 0 \\ CB & 0 & \dots & \dots & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ CA^{f-1}B & CA^{f-2}B & \dots & \dots & O \end{bmatrix}$$

Toeplitz matrices with  
Markov Parameters

$$H_s = \begin{bmatrix} I & 0 & \dots & 0 \\ CK & I & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ CA^{f-1}K & CA^{f-2}K & \dots & I \end{bmatrix}$$

“stochastic” ( $e$  as input)



# CCA, N4SID, MOESP

## 1. Using

State belongs to “past”

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

## 1. Using

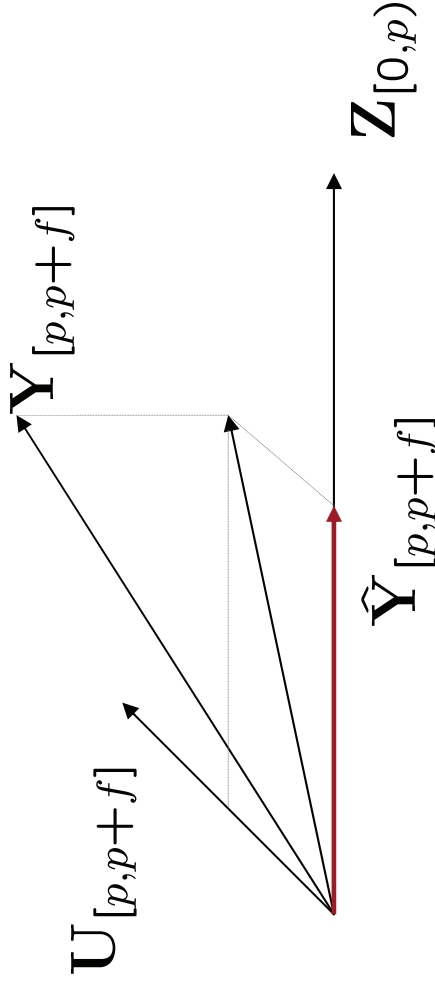
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$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

## 2.

Oblique projection (QR decomp.)

$$\Gamma \mathbf{X}_p \simeq \hat{\mathbf{Y}}_{[p,p+f]} := E_{||} \mathbf{U}_{[p,p+f]} \left[ \mathbf{Y}_{[p,p+f]} \mid \mathbf{Z}_{[0,p]} \right]$$





1. Using

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

State belongs to “past”

2.

$$\Gamma \mathbf{X}_p \simeq \hat{\mathbf{Y}}_{[p,p+f]} := \mathbf{E}_{||} \mathbf{U}_{[p,p+f]} \left[ \mathbf{Y}_{[p,p+f]} \mid \mathbf{Z}_{[0,p]} \right]$$

Oblique projection (QR decomp.)

3. Use “weighted SVD” to estimate  $\Gamma$  and/or  $\mathbf{X}_p$

$$U_n S_n V_n^\top \simeq W_r \hat{\mathbf{Y}}_{[p,p+f]} W_c$$

van Overschee – De Moor 1995



1. Using

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

State belongs to “past”

2.

$$\Gamma \mathbf{X}_p \simeq \hat{\mathbf{Y}}_{[p,p+f]} := E_{||\mathbf{U}_{[p,p+f]}} [\mathbf{Y}_{[p,p+f]} \mid \mathbf{Z}_{[0,p]}]$$

Oblique projection (QR decomp.)

3. Use “weighted SVD” to estimate  $\Gamma$  and/or  $\mathbf{X}_p$

$$U_n S_n V_n^T \simeq W_r \hat{\mathbf{Y}}_{[p,p+f]} W_c$$

van Overschee – De Moor 1995

4.

$$\hat{\Gamma} := W_r^{-1} U_n S_n^{1/2} \quad \hat{\mathbf{X}}_p := \hat{\Gamma}_r^{-L} \hat{\mathbf{Y}}_{[p,p+f]}$$





# PROBLEM: CLOSED-LOOP DATA

- **CCA, MOESP, N4SID** are not consistent if there is feedback [see e.g. Ljung and McKelvey 1996]
- What does go wrong?



# Where do things go wrong?

RECALL:

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

“ORTHOGONAL”

NOT “ORTHOGONAL”  
WITH FEEDBACK  
(future inputs depend upon past outputs/noise!)



# Where do things go wrong?

RECALL:

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

“ORTHOGONAL”

NOT “ORTHOGONAL”  
WITH FEEDBACK

(future inputs depend upon past outputs/noise!)

- Construction of the state! Need a **NEW LOOK** at stochastic realization theory when there are input (and possibly feedback) [see Chiuso and Picci 2002, Chiuso and Picci 2005]



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# Closed-loop (Subspace) ID

- **Early contributions** (Verhaegen '93/**Ljung-McKelvey'96**/ van Overschee-De Moor '97/Chou-Verhaegen '99/Katayama et al.)
- **Chiuso-Picci 2002/2003**: construction of the state with feedback-stochastic realization
- **Qin-Ljung 2003** Innovation Estimation Algorithm (IEM)
- **Jansson 2003** Subspace Algorithm with pre-estimation (SS-ARX) [similar to Peterzell 1996-Shi and McGregor 2001] [equivalent to Larimore 2004 (ADAPT<sub>x</sub>) (Chiuso ACC '07 to appear.)]
- **Chiuso-Picci 2004/2005** “Geometric version of SSARX” (PBSID) (more in the geometric spirit of subspace ID)

*Predictor Based Subspace Identification*

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# Assumptions

Most algorithms need the past horizon  $p$  to go to infinity with  $N$  while satisfying ( $\rho = \max |\lambda(\bar{A})|$ ):

$$p \geq \frac{\log N^{-d/2}}{\log |\rho|} \quad 1 < d < \infty$$
$$p = o((\log N)^\alpha) \quad \alpha < \infty$$



# Assumptions

Most algorithms need the past horizon  $p$  to go to infinity with  $N$  while satisfying ( $\rho = \max |\lambda(\bar{A})|$ ):

$$p \geq \frac{\log N^{-d/2}}{\log |\rho|} \quad 1 < d < \infty$$
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## Further Motivation:

*Subspace methods are “covariance based”:* it is well known that there is no efficient estimator for ARMA(X) processes based on a finite number of covariances [Walker ‘61, Porat-Friedlander ‘85] (unless the process is AR(X))



# Innovation Estimation

[Qin-Ljung 2003 (Chiuso-Picci 2002)]

RECALL:

LIES IN THE "PAST"

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

"DETERMINISTIC INPUT TERM"

IF WE KNEW  $\mathbf{E}_{[p,p+f]}$

$$\Gamma \mathbf{X}_p = E_{\|\mathbf{U}_{[p,p+f]} \vee \mathbf{E}_{[p,p+f]}} [\mathbf{Y}_{[p,p+f]} \mid \mathbf{Z}_{[0,p]}]$$

SIMILAR (BUT NOT THE SAME) Construction as Moonen et al (1988),  
(Willems 1986)

**INNOVATION ESTIMATION:** estimate innovation  $\hat{\mathbf{E}}_{[p,p+f]}$  (OLD CIRCLE OF IDEAS)





# “Future-Corrected” CCA

(ADAPT<sub>x</sub>) [Shi-McGregor ‘01 – Larimore ‘04]

RECALL:

$$\mathbf{Y}_{[p,p+f]} = \Gamma \mathbf{X}_p + H_d \mathbf{U}_{[p,p+f]} + H_s \mathbf{E}_{[p,p+f]}$$

IF WE KNEW  $\mathbf{H}_d$

ESTIMATE VIA ARX MODELING

$$\Gamma \mathbf{X}_p = E \left[ \underbrace{\mathbf{Y}_{[p,p+f]} - \hat{H}_d \mathbf{U}_{[p,p+f]}}_{\text{“Corrected” Future}} \mid \underbrace{\mathbf{Z}_{[0,p]}}_{\text{Past}} \right]$$

↖

Do **CCA** between corrected future and past! (closely related to **uCCA** in Peterzell et al. ‘96)



# Predictor-Based SID

[Jansson '03, Chiuso-Picci '04, Chiuso '05-'06]

- Based on “predictor model” as PEM identification.

$$\begin{cases} \mathbf{X}_{t+1} & = (A - KC)\mathbf{X}_t + B\mathbf{U}_t + K\mathbf{Y}_t \\ \mathbf{Y}_t & = C\mathbf{X}_t + \mathbf{E}_t \end{cases} \quad (1)$$

**Output of a system (known inputs) measured in white noise.**  
**CAREFUL: noise CORRELATED with future inputs (outputs)**

DISCUSSED IN DETAIL in CHIUSO-PICCI Automatica '05- IFAC '05



# NOTATION

(PREDICTOR MODEL  $\bar{A} := A - KC$ )

Observability matrix

$$\bar{F} := \begin{bmatrix} C \\ C\bar{A} \\ C\bar{A}^2 \\ \vdots \\ C\bar{A}^f \end{bmatrix}$$

``deterministic'' ( $u$  as input)

$$\bar{H}_d = \begin{bmatrix} 0 & 0 & 0 & \dots & 0 \\ CB & 0 & \dots & \dots & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ C\bar{A}^{f-1}B & C\bar{A}^{f-2}B & \dots & \dots & O \end{bmatrix}$$

Toeplitz matrices with  
Markov Parameters

$$\bar{H}_s = \begin{bmatrix} I & 0 & 0 & \dots & 0 \\ CK & I & \dots & \dots & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ C\bar{A}^{f-1}K & C\bar{A}^{f-2}K & \dots & \dots & I \end{bmatrix}$$

``stochastic'' ( $y$  as input)



# Predictor-Based SID

[Jansson '03, Chiuso-Picci '04, Chiuso '05-'06]

$$\mathbf{Y}_{[p,p+f]} = \bar{\Gamma} \mathbf{X}_p + \bar{H}_d \mathbf{U}_{[p,p+f]} + \bar{H}_s \mathbf{Y}_{[p,p+f]} + \mathbf{E}_{[p,p+f]}$$

1. Jansson 2003 (SSARX): Pre-estimates  $\hat{H}_d, \hat{H}_s$  using ARX and form

$$\mathbf{Y}_{[p,p+f]} - \hat{H}_d \mathbf{U}_{[p,p+f]} - \hat{H}_s \mathbf{Y}_{[p,p+f]} \simeq \bar{\Gamma} \mathbf{X}_p + \mathbf{E}_{[p,p+f]}$$

2. Chiuso-Picci 2004 (PBSID): Use a “geometric approach” (oblique projections) [equivalent to estimating  $f$  ARX models]



# “Unconstrained” least squares (PBSID)

$$\mathbf{Y}_{[p,p+f]} = \bar{\Gamma} \mathbf{X}_p + \bar{H}_d \mathbf{U}_{[p,p+f]} + \bar{H}_s \mathbf{Y}_{[p,p+f]} + \mathbf{E}_{[p,p+f]}$$

$$\begin{bmatrix} \mathbf{Y}_p \\ \mathbf{Y}_{p+1} \\ \vdots \\ \mathbf{Y}_{p+f} \end{bmatrix} = \begin{bmatrix} \Xi_0 \\ \Xi_1 \\ \vdots \\ \Xi_f \end{bmatrix} + \mathbf{Z}_{[0,p]} + \begin{bmatrix} 0 & 0 & \dots & 0 \\ \psi_{11} & 0 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ \psi_{ff} & \dots & \psi_{f1} & 0 \end{bmatrix} \mathbf{Z}_{[p,p+f]} + \begin{bmatrix} \mathbf{E}_p \\ \mathbf{E}_{p+1} \\ \vdots \\ \mathbf{E}_{p+f} \end{bmatrix}$$

$$\mathbf{E}_{[p,p+f]} := \begin{bmatrix} e_p \\ e_{p+1} \\ \vdots \\ e_{p+f} \\ e_{p+f+1} \\ \dots \\ e_{p+f+N-1} \end{bmatrix}$$

The vectorization has Singular Variance



# “Optimal” Predictor-Based SID

[PBSIDopt: Chiuso '05-'06]

Transform (quite naturally) the set of oblique projections (PBSID) in a weighted least-squares problem corresponding to minimizing  $\sum_{t=p}^{p+f+N-1} \|e_t\|^2$



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Transform (quite naturally) the set of oblique projections (PBSID) in a weighted least-squares problem corresponding to minimizing  $\sum_{t=p}^{p+f+N-1} \|e_t\|^2$

- Boils down to solving a (HUGE) equality constrained least squares problem (Computationally **very** intensive)
- The “optimized” PBSID automatically takes into account/enforces the
  - *block Hankel structure* of the data
  - *block Toeplitz structure* of the estimated parameters





# “Optimal” Predictor-Based SID

[PBSIDopt: Chiuso '05-'06]

- PBSIDopt  
(exactly!) equivalent to estimating a long VARX (of order  $p$ ) model

$$\mathbf{Y}_p = [ \hat{\Phi}_p \ \dots \ \hat{\Phi}_f \ \dots \ \hat{\Phi}_1 ] \mathbf{Z}_{[0,p]} + \mathbf{E}_p$$

followed by state estim. via SVD (= “model reduction”)



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followed by state estim. via SVD (= “model reduction”)

- Advantages for computational complexity (much easier)
- Take advantage of results from Stochastic Realization
- *Dahlen-Scherrer (2002)* proved that CCA is asymptotically equivalent to AR modeling + balanced model reduction
- Related approaches in [*Qin-Liung 2006, Onodera-Emoto-Qin 2006*]



# “Optimal” Predictor-Based SID

[PBSIDopt: Chiuso ‘05-’06]

$$\Phi_i := C \bar{A}^{i-1} [K \ B]$$

$$\hat{Y}_{[p,p+f]} = \begin{bmatrix} \hat{\Phi}_p & \dots & \hat{\Phi}_f & \dots & \hat{\Phi}_1 \\ 0 & \hat{\Phi}_p & \dots & \dots & \hat{\Phi}_2 \\ \vdots & \dots & \dots & \dots & \vdots \\ 0 & \dots & 0 & \hat{\Phi}_p & \dots \end{bmatrix} Z_{[0,p]}$$

Similar Ideas (Toeplitz structure) in *Peterzell et al. 1996* and perhaps relation with cut and shift (*Rapisarda-Willems*) **BUT** in our case the structure follows from an “optimally weighted” procedure, not just from imposing “ad-hoc” constraints



# Outline of the Talk

- Identification
  - Statement of the problem
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  - Main ideas
- Main (classical) Algorithms (CCA, N4SID, MOESP)
- Closed Loop Subspace Identification (New Alg.)
- **Recent contributions**
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  - Comparison of algorithms
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# Classical vs. New methods

- **CCA, N4SID, MOESP:**
  - **Consistent** (Open Loop) [*..many..*]
  - **Known variance formulas** [Bauer et al. 00, Jansson 00, Bauer-Ljung 02, Chiuso-Picci 04, Bauer 05]
  - **CCA efficient** (time-series, [Bauer05]) + **optimal** (white input, [Bauer-Ljung 02], *state joint model = state model of y given u* [Chiuso 07 TAC sub.])
- **SSARX** [Jansson 03], **IEM**. [Qin-Ljung 03], **PBSID** [Chiuso-Picci 04] **ADAPT<sub>x</sub>** [Larimore 04, ADAPT<sub>x</sub>]
  - **Consistent** (Closed Loop + Infinite `past`) [Chiuso-Picci, Automatica '05]
  - **Known Variance formulas** [Chiuso TAC'06]
  - **Similarities of PBSID with PEM** [Chiuso-Picci IFAC'05] **and relation SSARX/PBSID/CCA/FC-CCA/PBSIDopt** [TAC 07, Chiuso SYSID'06-Chiuso CDC'06-Chiuso AUTO 07, Chiuso ACC 07]
  - **Relation with other methods** (e.g. **IEM**)???



# Asymptotic Analysis

- $e(t)$  is a martingale difference sequence + constant conditional variance + bound on fourth order moments

Give two estimators  $\hat{\theta}_N, \hat{\psi}_N$  we shall use the notation

$$\hat{\theta}_N \doteq \hat{\psi}_N$$

if  $\hat{\theta}_N - \hat{\psi}_N = o_P(1/\sqrt{N})$  and hence  $\sqrt{N}\hat{\theta}_N$  and  $\sqrt{N}\hat{\psi}_N$  have the same asymptotic distribution



# Asymptotic Variance

- Open Loop (N4SID, CCA, MOESP)  
many papers:
  - Bauer-Deistler-Scherrer (99), Bauer-Jansson (00), Jansson (00), Bauer-Ljung (02), Bauer (05), Chiuso-Picci (04) [*+ Lovera-Bittanti using Bootstrap (00)*]
- Closed Loop: IEM, SSARX, PBSID (WFA)
  - Chiuso (TAC 2006)



# Asymptotic Variance (PBSID) (Chiuso, TAC06)

Let  $\tilde{A}_N := \hat{A}_N - A_N$  ( $\tilde{B}_N$  etc.), then  $\sqrt{N}\text{vec}(\tilde{A}_N)$ ,  $\sqrt{N}\text{vec}(\tilde{B}_N)$  and  $\sqrt{N}\text{vec}(\tilde{C}_N)$  are jointly asymptotically normal with as. variance matrix

$$\begin{pmatrix} M_{A1}P & M_{A2}\bar{P} \\ M_{B1}P & M_{B2}\bar{P} \\ M_C P & 0 \end{pmatrix} \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{11}^T & \Sigma_{22} \end{pmatrix} \begin{pmatrix} M_{A1}P & M_{A2}\bar{P} \\ M_{B1}P & M_{B2}\bar{P} \\ M_C P & 0 \end{pmatrix}^T$$

1.  $P$  and  $\bar{P}$  permutation matrices
2.  $\Sigma_{ij}$  input-output conditional covariance matrices
3.  $M_{A1}$ ,  $M_{A2}$ , depends on the “true” system parameters and some cond. cov. matrices (state-input-output)





# Practical Problem

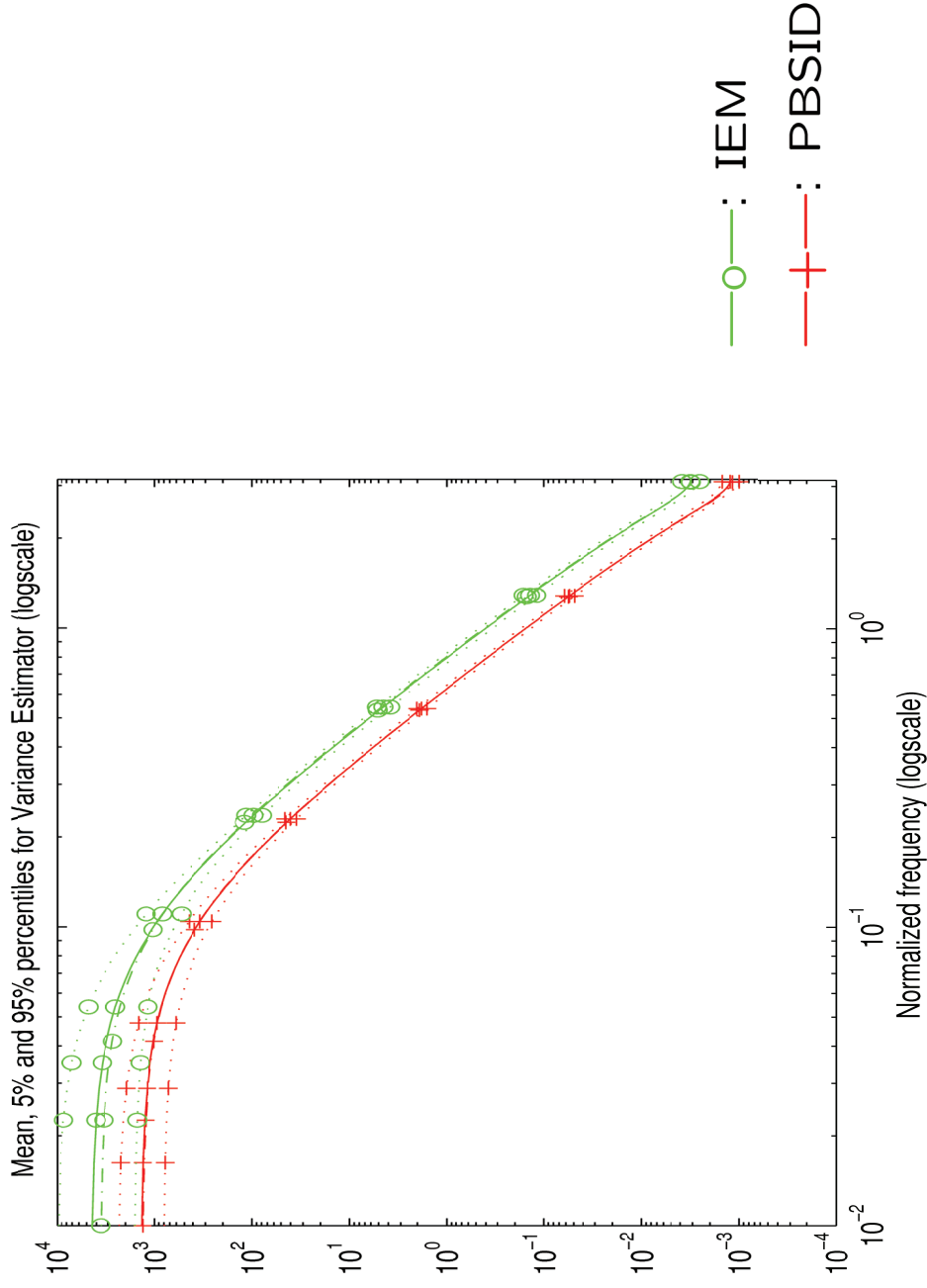
- Want to assess model accuracy starting from data alone.
- Do the ASVAR formulas help us to this purpose?
  - Compute variance from data:
    - Difficult/Computationally expensive?
    - Reliable?

**SIMPLE (O(N)) ALGORITHM TO APPROXIMATE VARIANCE FROM DATA**



# Estimated vs. Asympt. Variance

Order 1

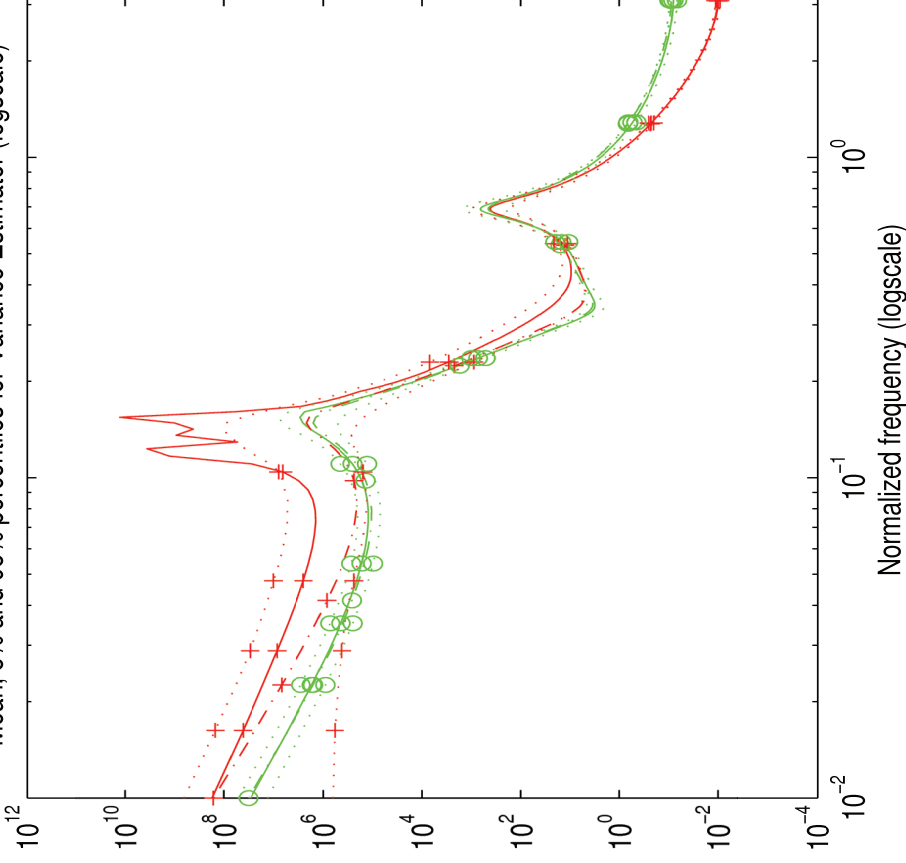




# Estimated vs. Asympt. Variance

Order 5

Mean, 5% and 95% percentiles for Variance Estimator (logscale)



—○—: IEM

—+—: PBSID



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# PBSID/PBSIDopt vs. CCA

(open loop)

- PBSID asympt. Eq. to CCA with white inputs
- PBSIDopt no worse than CCA

[Chiuso TAC 07, to appear]

**Theorem.** Denote with  $\hat{\Theta}^{CCA}$  and  $\hat{\Theta}^{P_{opt}}$  the estimators of any system invariant using respectively CCA and PBSID<sub>opt</sub>; then,

$$\text{AsVar} \{ \sqrt{N} \hat{\Theta}^{P_{opt}} \} \leq \text{AsVar} \{ \sqrt{N} \hat{\Theta}^{CCA} \}.$$



# SSARX vs. PBSID vs. ADAPT<sub>x</sub> + Ljung-McKelvey '96 (closed loop)

- SSARX, ADAPT<sub>x</sub> and PBSID are asymptotically equivalent

In the sense of

VanOverschee-DeMoor '95

- PBSIDopt is a “weighted version” of Ljung-McKelvey 1996

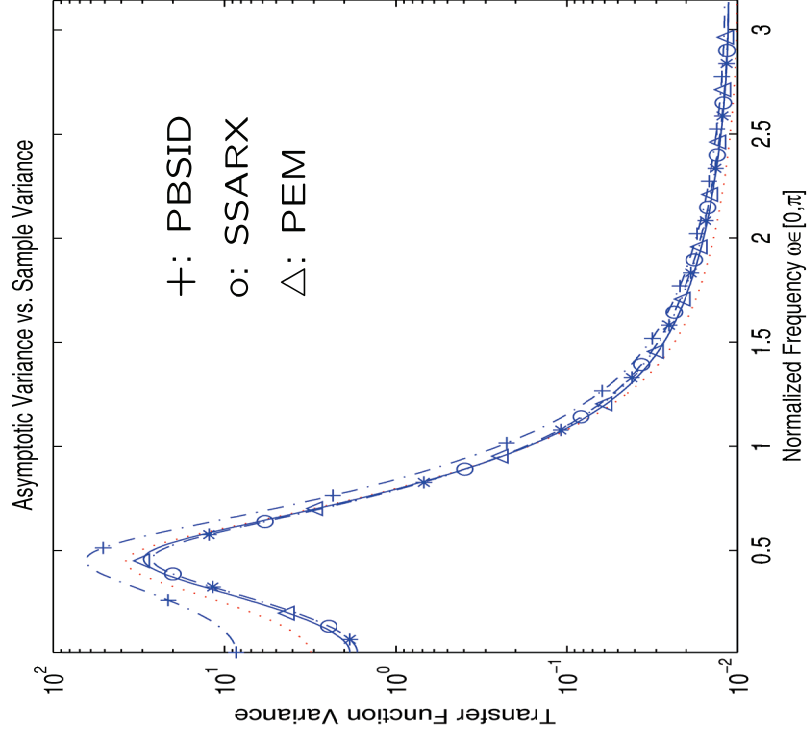
*[Chiuso, Automatica 2007 + ACC 07 (to appear)]*



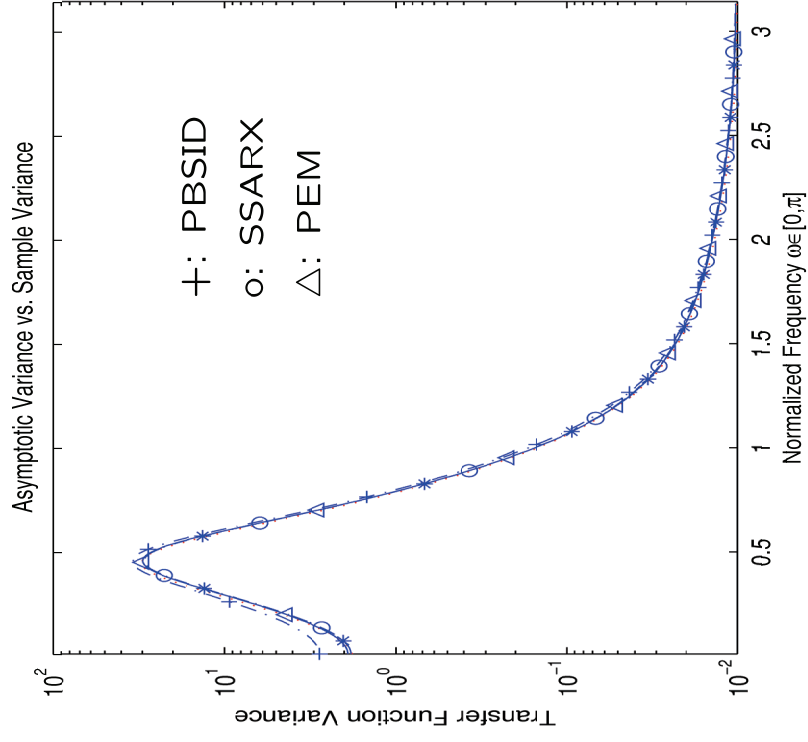
# SIMULATIONS

(PBSID vs. SSARX: II order):

$f = 10, p = 10, N = 1000$

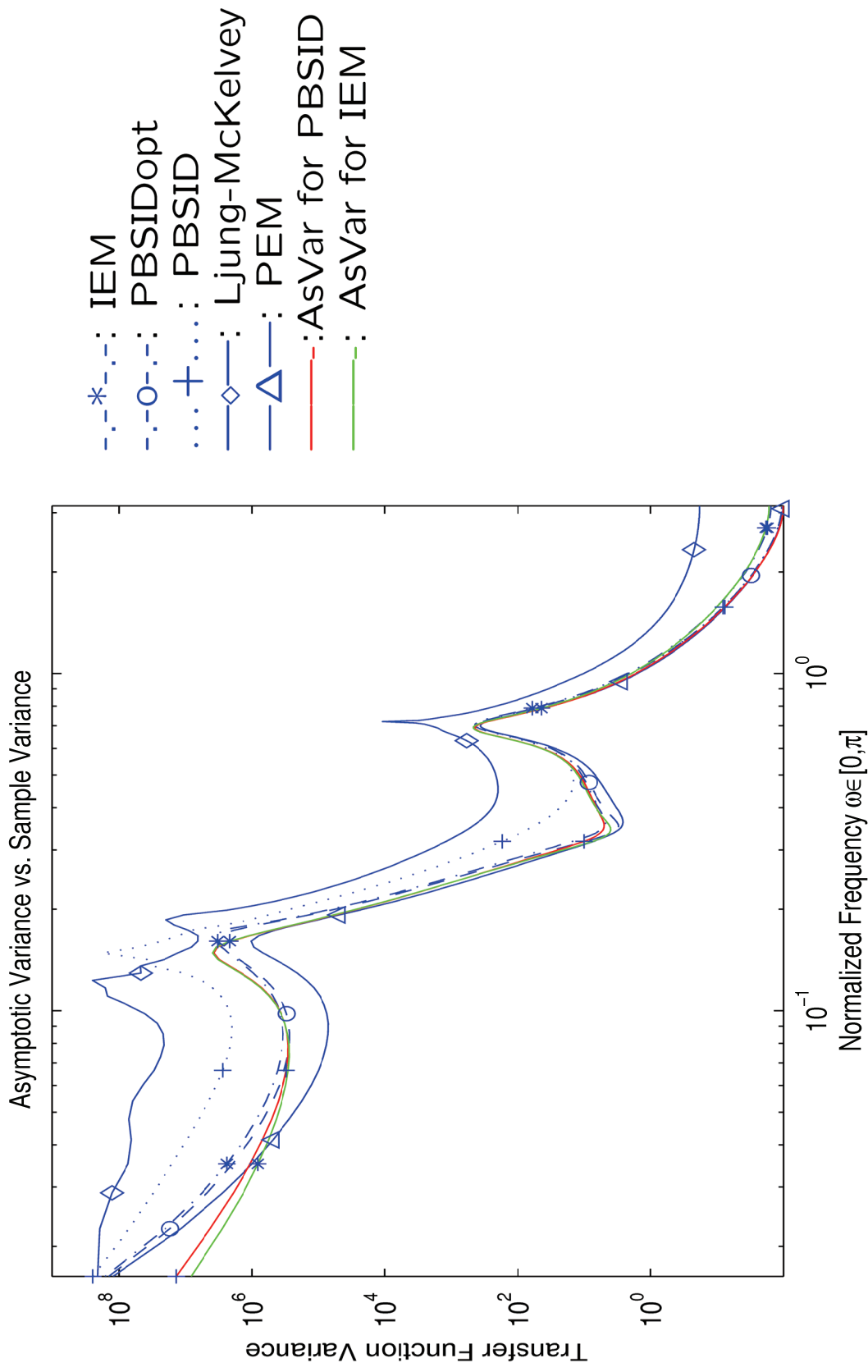


$f = 10, p = 30, N = 3000$





# Fifth-order example (CL)







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# Role of $f$ (example)

- MODEL
$$y(t) - 0.5y(t - 1) = u(t - 1) + e(t) + 0.5e(t - 1)$$
- INPUT  $u(t)$ 
$$u(t) - 0.9u(t) = n(t) - \alpha n(t - 1)$$

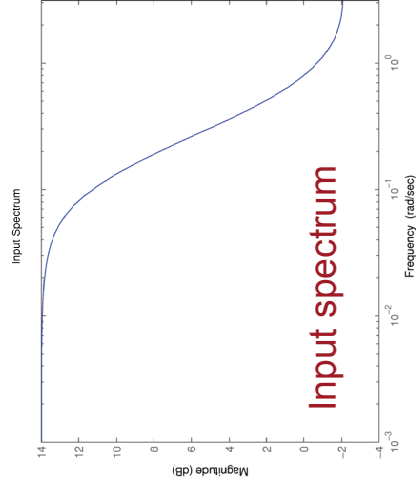
$n(t)$  white noise
- “Efficiency” index

$$Eff(f) := \frac{\int_0^{2\pi} \text{AsVar} \{ \hat{F}_f(e^{j\omega}) \} d\omega}{\int_0^{2\pi} \text{CRLB}_{\hat{F}}(j\omega) d\omega} \quad (1)$$

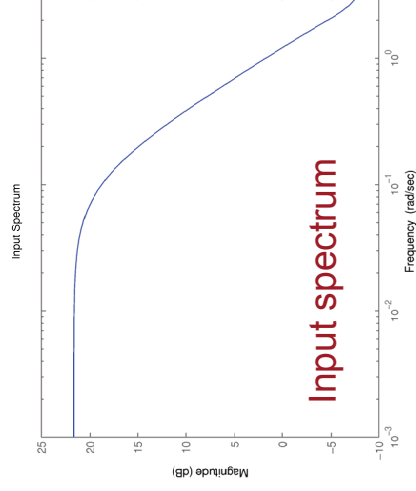


# Role of $f$

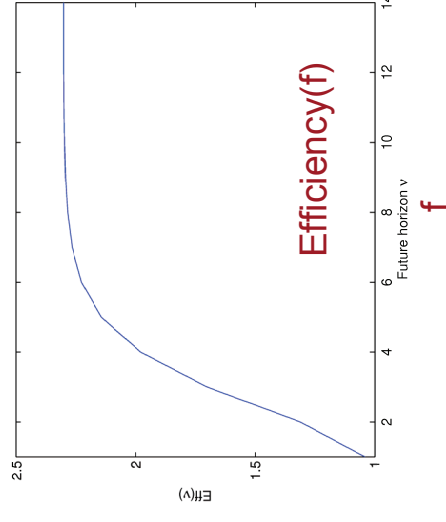
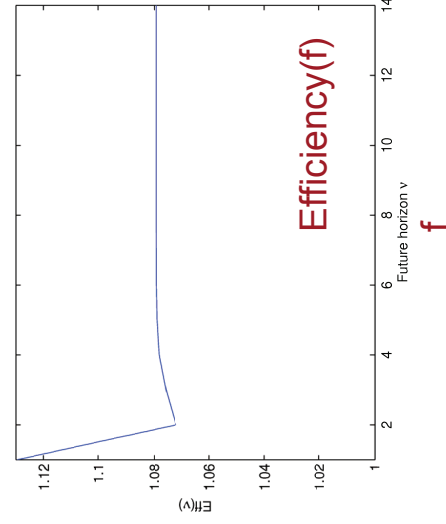
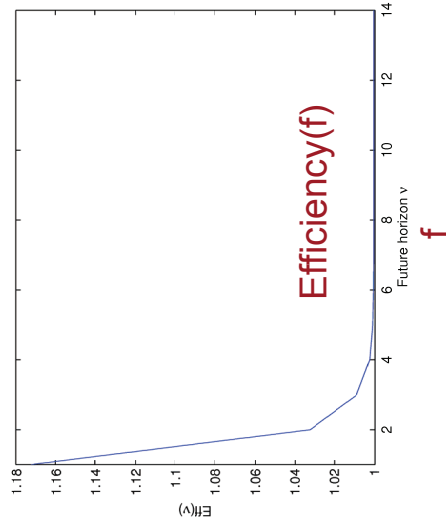
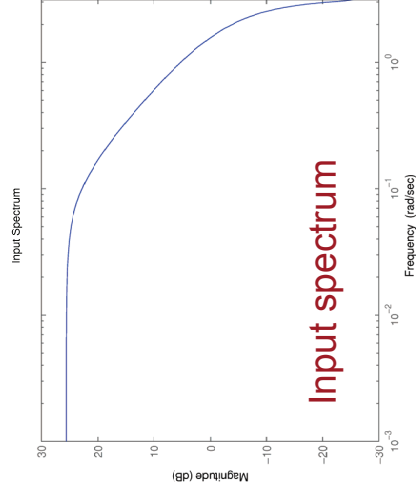
$$\alpha = 0.5$$



$$\alpha = -0.2$$



$$\alpha = -0.9$$



Stockholm - May 11th, 2007



# Role of $f$ (future horizon)

Can show that the estimated state via SSARX-ADAPT-X-PBSID is of the form

$$\hat{X}_t - X_t = S_n^{-1}(f) \hat{E} [V_t(f) | Z_{[t_0, t]}] + \\ - S_n^{-1}(f) U_n^\top(f) W_L^{-1}(f) \tilde{H}_f^d \hat{E} [U_{[t, T]} | Z_{[t_0, t]}]$$



# Role of $f$ (future horizon)

Fixed basis

Can show that the estimated state via SSARX-ADAPT-X-PBSID is of the form

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$$-S_n^{-1}(f) U_n^T(f) W_L^{-1}(f) \tilde{H}_f^d \hat{E} [U_{[t,T]} | Z_{[t_0,t]}]$$

Goes to zero with  $N$ , Variance independent of  $f$  for  $u$  white

Error in estimation of the Markov Parameters



# Role of $f$ (future horizon)

Fixed basis

Can show that the estimated state via SSARX-ADAPT-X-PBSID is of the form

$$\hat{X}_t - X_t = S_n^{-1}(f) \tilde{E} [V_t(f) | Z_{[t_0,t]}] +$$

$$-S_n^{-1}(f) U_n^T(f) W_L^{-1}(f) \tilde{H}_f^d \tilde{E} [U_{[t,T]} | Z_{[t_0,t]}]$$

Goes to zero with  $N$ , Variance independent of  $f$

Decreasing with  $f$

Error in estimation of the Markov Parameters

Increases with  $f$ , unless  $u(t)$  white, in which case vanishes with  $N$

- For white inputs variance is monotonically decreasing with  $f$  [Bauer-Ljung] (the second term on the right hand side can be neglected while the first is monotonically decreasing)



# Role of $f$ (future horizon)

## Theorem [Chiuso, TAC 2007, Submitted]

Assume the input process is such that the rows of  $X_t^u$  are, up to  $o_P(1/\sqrt{N})$  terms, linear combinations of the rows of  $X_t$  (or equivalently  $\text{span}\{\mathbf{x}^u(t)\} =: \mathcal{X}_t^u \subseteq \mathcal{X}_t := \text{span}\{\mathbf{x}(t)\}$ ); then the asymptotic variance of any system invariant estimated using the FC-CCA algorithm is a monotonically non increasing function of  $\nu$ .



# Role of $f$ (future horizon)

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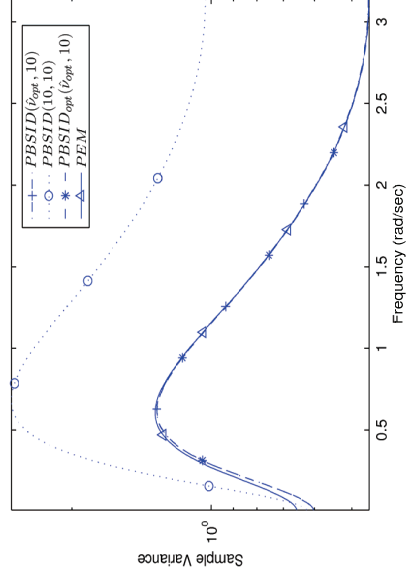
**REMARK:** FC-CCA as. eq. to PBSID, SSARX



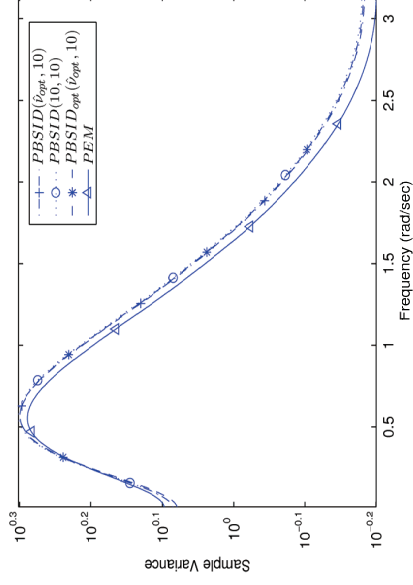


# Estimation of the “optimal” $f$ using asymptotic variance formulas

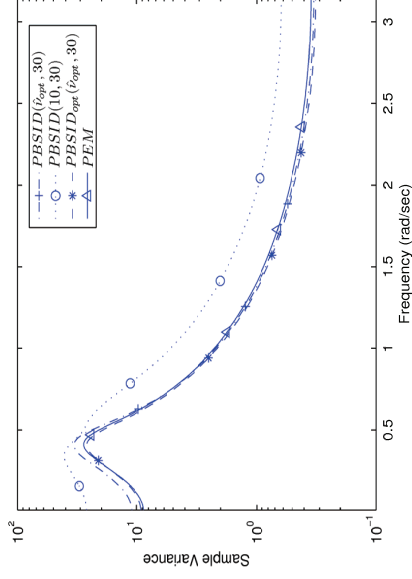
$$\alpha = -0.9$$



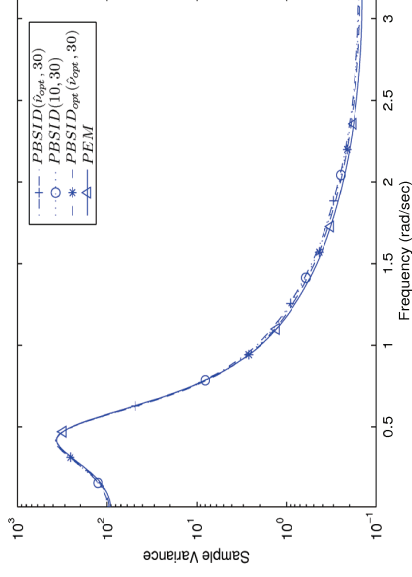
$$\alpha = -0.2$$



Sensitive to  $\nu$



Not sensitive to  $\nu$





# Remark

- It was claimed in [Larimore 04] that the ADAPT<sub>x</sub> software is efficient as  $p$  and  $f$  go to infinity with the number of data.
- The plots in the previous slides show that indeed this is not true.
- Hence the quest for an efficient subspace procedure (if any) is still open.



# CONCLUSIONS

- SUBSPACE ID is (ALMOST) a jungle of algorithms  
**BUT**  
Stochastic realization theory provides a unifying framework which has proven useful in deriving/comparing algorithms.
- Statistical analysis has been developed a lot since the late 90's. (Variance computation is less intensive than algorithm itself).  
Variance formulas provide also a natural tool to compare algorithms directly on data  
(we have used it to select the ``optimal'' f)



# OPEN QUESTIONS!

- Statistically efficient procedures (*WHERE DOES STOCH. REALIZATION FAIL?*)
  - Choice of  $f$ ?
  - Undermodeling?
- Performance with finite data (comparison, when available, holds only asymptotically)
  - *Quality of the asymptotic approximation as a function of location in parameter space*
- Relation with ARX + model reduction
- Enforcing structure (*important for MIMO application when number of inputs and output is very large (50-100!)*)...





# Main References

CHECK OUT:

[www.dei.unipd.it/~chiuso](http://www.dei.unipd.it/~chiuso)

# THANK YOU